

# LDA In Slow Motion - Journey of a Parametric Topic Model

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# What is a topic?

- ▶ "I call the same thing *element* and *topos*; for an element or a topos is a heading under which many enthymemes fall." - Aristotle
- ▶ An **enthymeme**, in its modern sense, is an informally stated syllogism (a three-part deductive argument) with an unstated assumption that must be true for the premises to lead to the conclusion
  - ▶ Aristotle. *Topica*. Translated by E. S. Forster. Loeb Classical Library. Cambridge: Harvard University Press, 1989
    - [en.wikipedia.org/wiki/Topics\\_\(Aristotle\)](https://en.wikipedia.org/wiki/Topics_(Aristotle))
- ▶ The Topics contains and relies upon Aristotle's definition of reasoning; *a verbal expression in which, certain things having been laid down, other things necessarily follow from these*

# What is a topic? – Aristotle's view

## Case 1:

- ▶ Major premise: All men are mortal
- ▶ Minor premise: All Greeks are men
- ▶ Conclusion: All Greeks are mortal

## Case 2:

- ▶ Major premise: No homework is fun
- ▶ Minor premise: Some reading is homework
- ▶ Conclusion: Some reading is not fun

# What is a topic? – LDA's view

## Case 1:

- ▶ men, mortal co-occur
- ▶ Greeks, men co-occur
- ▶ Conclusion: men, mortal and Greeks must be related under some notion

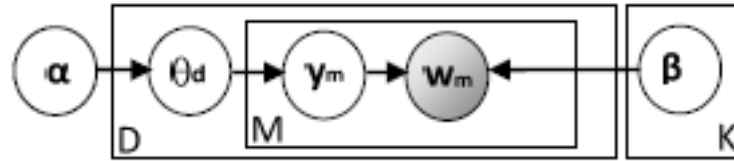
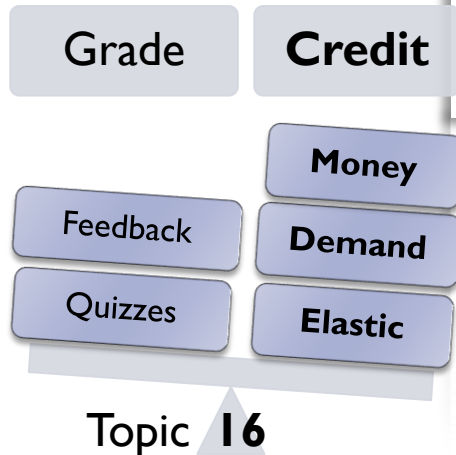
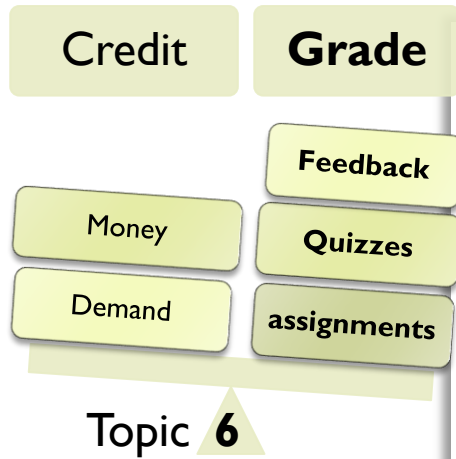
## Case 2:

- ▶ homework, fun co-occur
- ▶ reading, homework co-occur
- ▶ Conclusion: homework, fun and reading must be related under some *other* notion – not the one involving Greeks and mortal

# What is a topic? – LDA's view

- ▶ Repeated syllogisms referring to the same “theme” introduces word co-occurrence
- ▶ That words under a “theme” are related is a direct outcome of the way text presents itself in English
- ▶ We will hardly see a syllogism like
  - ▶ Major premise: All men are mortal
  - ▶ Minor premise: Some reading is homework
  - ▶ Lack of Conclusion: All men are mortal AND Some reading is homework? ( **When I write my first draft of a paper after prolonged coding, it reads like this – all I can think of is “Our model performs better...” and “for (k=0; k < ...)”** )

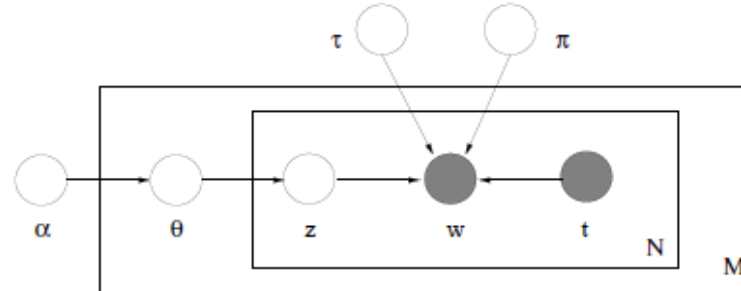
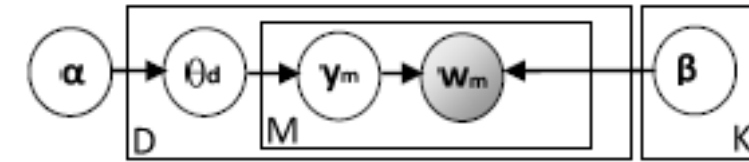
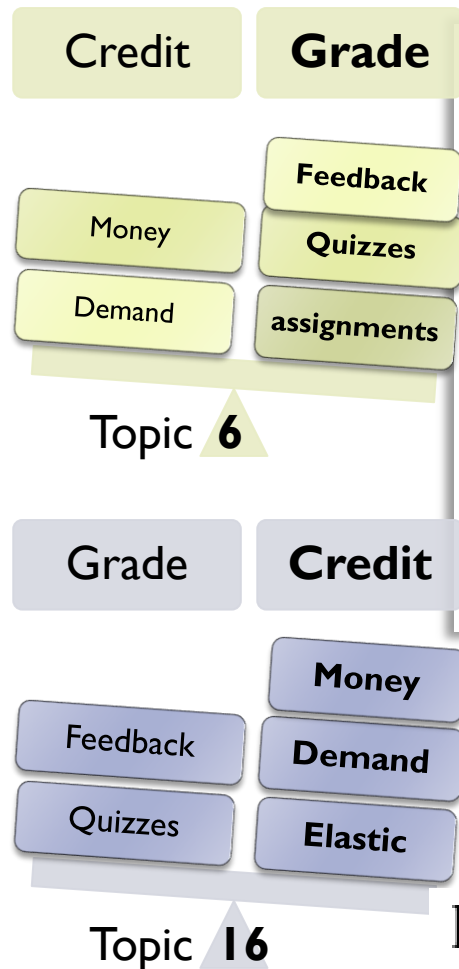
# LDA



$\beta_{LDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_k$
$\wedge$				
<b>K</b>				
$\vee$				

- $D$  = number of documents
- $M$  = number of word positions in the document
- $K$  = number of latent dimensions a.k.a topics/clusters
- $V$  = total number of words in corpus
- $w_m$  = observed word in document  $d \in D$
- $y_m$  = hidden topic  $i \in K$  for word  $w_m$  in document  $d$
- $\theta_d$  = document  $d$ 's hidden distribution over  $K$  topics
  - These 3 grow with the data (think about  $D$ )
- $\alpha$  = parameter of LDA
- $\beta$  = parameter of LDA (that we finally like to visualize)
  - These 2 does not grow with data (independent of  $D$ )

# LDA and TagLDA



$\beta_{LDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_k$
$\wedge$				$ $
<b>K</b>				$ $
$\vee$				$ $

$\tau_{TagLDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_k$
$\wedge$				$\neq  $
<b>K</b>				$\neq  $
$\vee$				$\neq  $

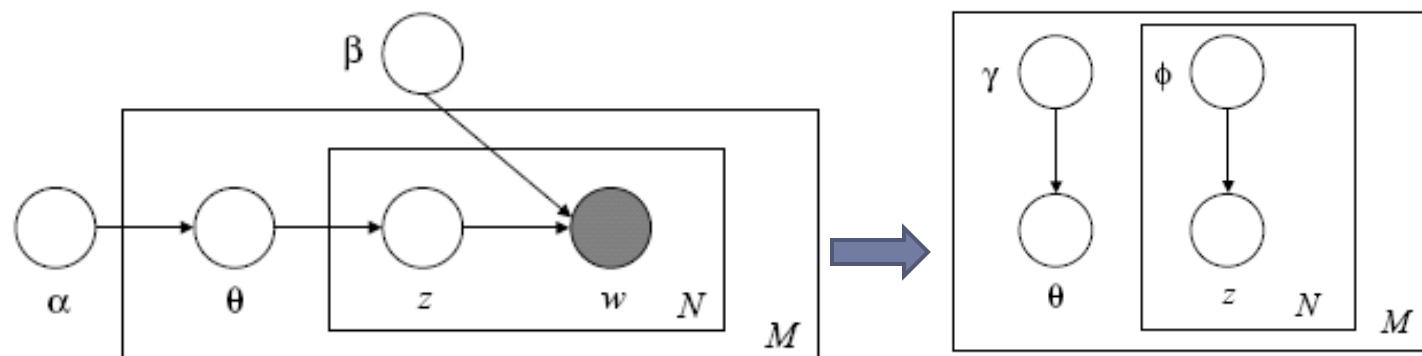
$\pi_{TagLDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_t$
$\wedge$				$\neq  $
<b>T</b>				$\neq  $
$\vee$				$\neq  $

Figure 1: Graphical model representations of LDA and TagLDA models

# Parameters of LDA

- ▶  $\beta$  is a set of distributions (topics)
- ▶ For each dataset, we only see its values for the words present in the dataset
- ▶  $\beta$  is row normalized and is represented in log-space since we calculate log likelihoods
- ▶  $\alpha_i$  is a prior for  $\theta_i$  encoding the belief that a document is “about a topic  $i$ .” In basic LDA, all  $\alpha_i$ ’s are equal to  $\alpha$  (symmetric)
  - ▶ However,  $\beta$  needs to be updated as each document is encountered
  - ▶  $\beta_{\text{sufficient-statistics}}$  is a SINGLE auxiliary storage for  $\beta$  that contains un-normalized entries for  $\beta$
  - ▶  $\beta_{\text{sufficient-statistics}}$  has the same dimensions as  $\beta$
  - ▶  $\alpha_{\text{sufficient-statistics}}$  similarly is an auxiliary storage for  $\alpha$

# Variational Inference in LDA

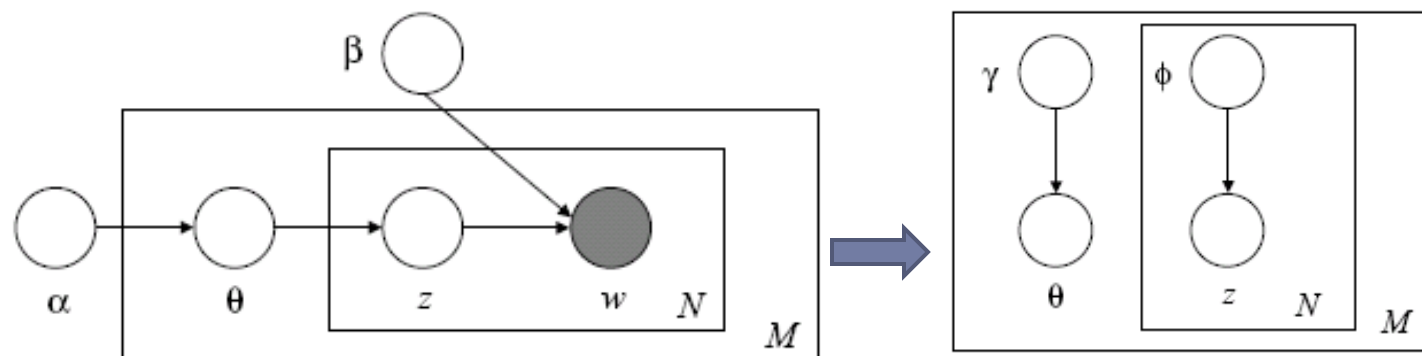


$$p(\theta, z | \mathbf{w}, \alpha, \beta) = \frac{p(\theta, z, \mathbf{w} | \alpha, \beta)}{p(\mathbf{w} | \alpha, \beta)}$$

$$p(\mathbf{w} | \alpha, \beta) = \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \int \left( \prod_{i=1}^k \theta_i^{\alpha_i - 1} \right) \left( \prod_{n=1}^N \sum_{i=1}^k \prod_{j=1}^V (\theta_i \beta_{ij})^{w_n^j} \right) d\theta,$$

- ▶ Normalizing constant is intractable due to the coupling between  $\theta$  and  $\beta$  in the summation over latent topics
- ▶ Key motivation - compute the posterior distribution of the hidden variables ( $z$  and  $\theta$ ) given a document (word and document distribution over topics)

# Variational Inference in LDA



- ▶ Find tightest possible lower bound to the value of the expression, by using optimization of variational parameters
  - ▶ A simple way to obtain a tractable family of lower bounds is to consider *simple modifications* of the original graphical model in which some of the edges and nodes are removed
- ▶ In particular, in the LDA model shown above, the coupling between  $\theta$  and  $\beta$  arises due to the edges between  $\theta$ ,  $z$ , and  $w$ 
  - ▶ By dropping *these edges* and the *w node*, and using free variational parameters in the resulting simplified graphical model, we obtain a *family* of distributions ( for  $\phi$  and  $\gamma$ ) on the latent variables  $\theta$  and  $z$

# General EM algorithm

- ▶ **Local** optimization algorithm
- ▶ Given a set of observed (visible) variables  $V$ , a set of unobserved (hidden/ latent/missing) variables  $H$ , and model parameters  $\theta$ , optimize the log probability of the observed data

$$\mathcal{L}(\theta) = \log p(V|\theta) = \log \int dH p(H, V|\theta)$$

- ▶ Using Jensen's inequality for any distribution of hidden states  $q(H)$  [Note: log is concave]

$$\mathcal{L}(\theta) = \log \int dH q(H) \frac{p(H, V|\theta)}{q(H)} \geq \int dH q(H) \log \frac{p(H, V|\theta)}{q(H)} = F(q, \theta)$$

- ▶ Normally we are given only  $V$ , not  $(H, V)$
- ▶ In the EM algorithm, we alternately optimize  $F(q, \theta)$  w.r.t  $q$  and  $\theta$

Acknowledgements: Matthew Beal's lecture notes

# General EM algorithm

- ▶ The lower bound on the log likelihood

$$\mathcal{F}(q, \theta) = \int dH q(H) \log \frac{p(H, V | \theta)}{q(H)} = \int dH q(H) \log p(H, V | \theta) + \mathcal{H}(q)$$

- ▶ E step: optimize  $\mathcal{F}(q, \theta)$  w.r.t the distribution over hidden variables given the parameters

$$q^{(k)}(H) := \operatorname{argmax}_{q(H)} \mathcal{F}(q(H), \theta^{(k-1)})$$

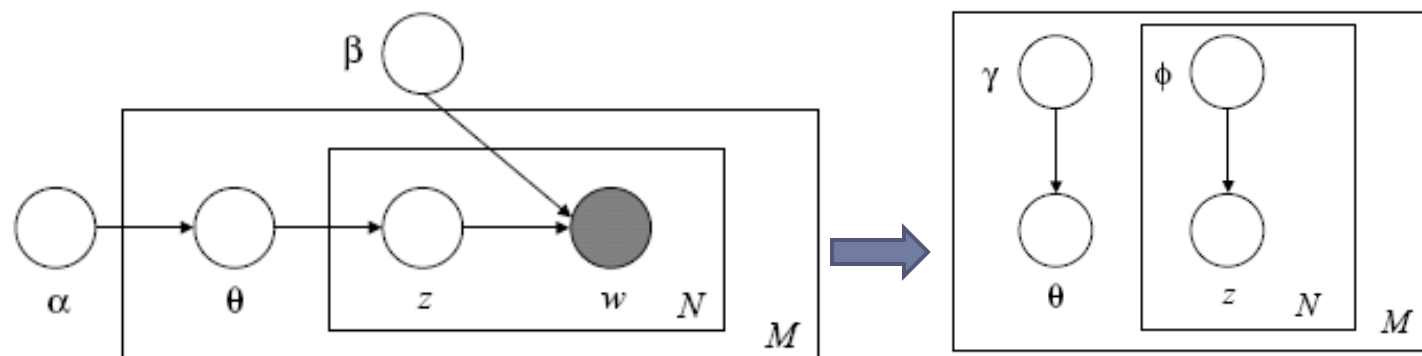
- ▶ M step: maximize  $\mathcal{F}(q, \theta)$  w.r.t the parameters given the hidden distribution

$$\theta^{(k)} := \operatorname{argmax}_{\theta} \mathcal{F}(q^{(k)}(H), \theta) = \operatorname{argmax}_{\theta} \int dH q^{(k)}(H) \log p(H, V | \theta)$$

- ▶ **Take Home Message:**

- ▶ E step: fill in values for the hidden variables according to their posterior probabilities
- ▶ M step: learn model as if hidden variables were not hidden

# Variational Inference (VBEM) in LDA



- ▶ Note:  $\gamma$  (Dirichlet) and  $(\phi_1, \phi_2, \dots, \phi_N)$  (Multinomials) are *free parameters*

$$q(\theta, \mathbf{z} | \gamma, \phi) = q(\theta | \gamma) \prod_{n=1}^N q(z_n | \phi_n)$$

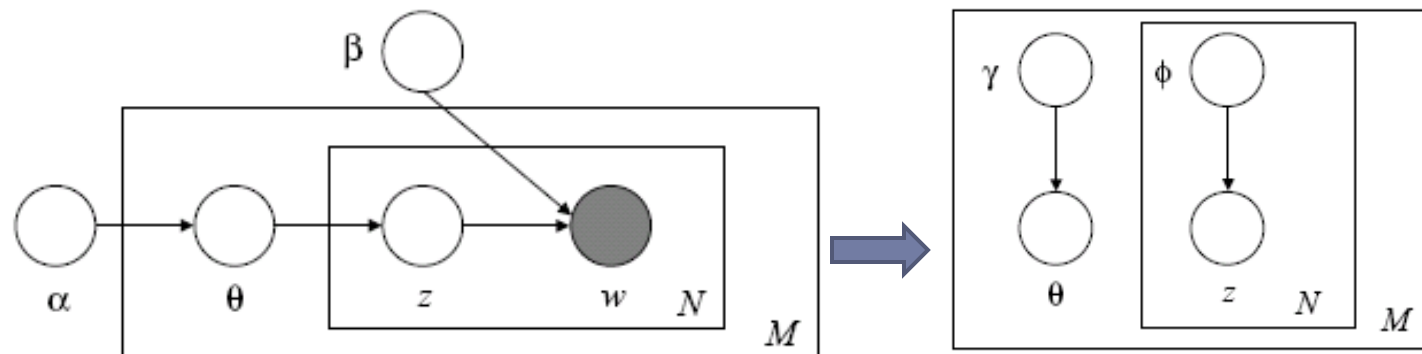
$$(\gamma^*, \phi^*) = \underset{(\gamma, \phi)}{\operatorname{argmin}} D(q(\theta, \mathbf{z} | \gamma, \phi) || p(\theta, \mathbf{z} | \mathbf{w}, \alpha, \beta))$$

$$\phi_{ni} \propto \beta_{iw_n} \exp\{E_q[\log(\theta_i) | \gamma]\}$$

$$\gamma_i = \alpha_i + \sum_{n=1}^N \phi_{ni}$$

$$E_q[\log(\theta_i) | \gamma] = \Psi(\gamma_i) - \Psi(\sum_{j=1}^k \gamma_j)$$

# Variational Inference (VBEM) in LDA



$$\beta_{ij} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi^*_{dni} w^j_{dn}$$

$$\alpha_{\text{new}} = \alpha_{\text{old}} - H(\alpha_{\text{old}})^{-1} g(\alpha_{\text{old}}) \text{ See original paper for } \alpha \text{ optimization}$$

## Take Home Message:

- ▶ Training: Initialize  $\alpha$  and  $\beta$  to initial values  $\rightarrow$  compute  $\gamma, \phi$  from observed documents  $\rightarrow$  again compute  $\alpha$  and  $\beta$  from  $\gamma, \phi$
- ▶ Testing: Compute  $\gamma, \phi$  for unseen documents from learned  $\alpha$  and  $\beta$  for training documents. Update  $\alpha$  and  $\beta$  accordingly
  - ▶ Note that testing is implicit in training (why?)



# Intuitions for VBEM in LDA

- ▶ Rarely, if ever, does a document contain all words belonging to the same topic
  - ▶ Represent a document as a distribution over topics
- ▶ Given any decently sized document set, it is not possible to exactly estimate the values of the distribution over all combinations of observations in our lifetime
- ▶ Assume simpler distributions that tally with our beliefs and compute **tractable expectations rather than exact estimations**
- ▶ Find as tight a lower bound to the absolute truth – the probability of the data given the model

# Expectation Maximization (EM) for LDA

```
▶ compute_initial_guesses() // for beta (corpus distribution)
prev_ELBO = 0;
while (converged > tolerance)
{
    e-step(); // zero initialize sufficient statistics; initialize and update
per document distributions from corpus distributions; sufficient
statistics depend on per document distributions
    m-step(); // update corpus distributions from sufficient statistics
    curr_ELBO = compute_likelihood(); // sums up the individual
doc_likelihoods from compute_doc_likelihood()s in doc-e-step()s
    converged = (prev_ELBO-curr_ELBO)/prev_ELBO;
    prev_ELBO = curr_ELBO;
}
```

# Computing Likelihood for LDA

## ► Variational Inference

- Recall the variational distribution we had before

$$q(\theta, \mathbf{z} | \gamma, \phi) = q(\theta | \gamma) \prod_{n=1}^N q(z_n | \phi_n)$$

$$\begin{aligned} \log p(\mathbf{w} | \alpha, \beta) &= \log \int \sum_{\mathbf{z}} p(\theta, \mathbf{z}, \mathbf{w} | \alpha, \beta) d\theta \\ &= \log \int \sum_{\mathbf{z}} \frac{p(\theta, \mathbf{z}, \mathbf{w} | \alpha, \beta) q(\theta, \mathbf{z})}{q(\theta, \mathbf{z})} d\theta \\ &\geq \int \sum_{\mathbf{z}} q(\theta, \mathbf{z}) \log p(\theta, \mathbf{z}, \mathbf{w} | \alpha, \beta) d\theta - \int \sum_{\mathbf{z}} q(\theta, \mathbf{z}) \log q(\theta, \mathbf{z}) d\theta \\ &= E_q[\log p(\theta, \mathbf{z}, \mathbf{w} | \alpha, \beta)] - E_q[\log q(\theta, \mathbf{z})] \end{aligned}$$

$$\log p(\mathbf{w} | \alpha, \beta) = L(\gamma, \phi; \alpha, \beta) + D(q(\theta, \mathbf{z} | \gamma, \phi) || p(\theta, \mathbf{z} | \mathbf{w}, \alpha, \beta))$$

Where,  $L(\gamma, \phi; \alpha, \beta) = E_q[\log p(\theta, \mathbf{z}, \mathbf{w} | \alpha, \beta)] - E_q[\log q(\theta, \mathbf{z})]$

# Computing Likelihood for LDA

## ► Variational Inference

- We now expand the  $L(\gamma, \phi; \alpha, \beta)$  into 5 core expressions

$$L(\gamma, \phi; \alpha, \beta) = E_q[\log p(\theta | \alpha)] + E_q[\log p(z | \theta)] + E_q[\log p(\mathbf{w} | z, \beta)] \\ - E_q[\log q(\theta)] - E_q[\log q(z)]$$

- We expand the expression in terms of the model parameters  $(\alpha, \beta)$  and the variational parameters  $(\gamma, \phi)$

compute\_doc\_likelihood()

$$L(\gamma, \phi; \alpha, \beta) = \log \Gamma\left(\sum_{j=1}^k \alpha_j\right) - \sum_{i=1}^k \log \Gamma(\alpha_i) + \sum_{i=1}^k (\alpha_i - 1)(\psi(\gamma_i) - \psi(\sum_{j=1}^k \gamma_j)) \\ + \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} (\psi(\gamma_i) - \psi(\sum_{j=1}^k \gamma_j)) \\ + \sum_{n=1}^N \sum_{i=1}^k \sum_{j=1}^V \phi_{ni} w_n^j \log \beta_{ij} \\ - \log \Gamma\left(\sum_{j=1}^k \gamma_j\right) - \sum_{i=1}^k \log \Gamma(\gamma_i) + \sum_{i=1}^k (\gamma_i - 1)(\psi(\gamma_i) - \psi(\sum_{j=1}^k \gamma_j)) \\ - \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} \log \phi_{ni}$$

# Expectation Maximization (EM) for LDA

doc-e-step() {  $\alpha^{(suffstats)} = 0; \beta^{(suffstats)} = 0$

initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$

initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$

prev\_ELBO = 0; t = 0;

repeat

for  $n = 1$  to  $N$  {

for  $i = 1$  to  $K$

$$\phi_{n,i}^{t+1} = \beta_{i,w_n} \exp(\psi(\gamma_i^t) - \psi(\sum_{j=1}^K \gamma_j^t));$$

normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1

$$\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$$

curr\_ELBO = compute\_doc\_likelihood();

converged = (prev\_ELBO - curr\_ELBO)/prev\_ELBO;

until converged > tolerance or t < Max\_T for e-step

for  $n = 1$  to  $N$

for  $i = 1$  to  $K$

$$\beta_{i,w_{d,n}}^{(suffstats)} += \phi_{n,i} w_{d,n};$$

for  $i = 1$  to  $K$

$$\alpha^{(suffstats)} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$$

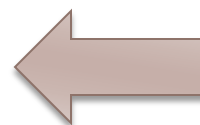
}

This expression is not needed since it is same for all  $n$  in  $N_d$



Store final value of curr\_ELBO as the doc\_likelihood for this document

Precompute this term in the alpha update first



# Synthetic Example for LDA

- ▶ `model.input =`
  - ▶ 5 0:1:1 1:1:1 1:1:1 3:1:1 3:1:1 //wordID:count:tagID (count is always 1)
  - ▶ 6 1:1:1 2:1:1 2:1:1 2:1:1 3:1:1 3:1:1 // tagID not used in LDA but
  - ▶ 4 4:1:1 5:1:1 6:1:1 0:1:1 // used in TagLDA (Later slides)
  - ▶ 6 4:1:1 5:1:1 4:1:1 5:1:1 4:1:1 5:1:1 // 6 = num. of positions
- ▶ `term_to_id.txt=`
  - ▶ Gilchrist=0 wicket=1 keeper=3 captain=2  
Jones=4 data=5 architect=6
- ▶ `model.input + term_to_id.txt =`
  - ▶ doc0: Gilchrist wicket wicket keeper keeper
  - ▶ doc1: wicket captain captain captain keeper keeper
  - ▶ doc2: Jones data architect Gilchrist
  - ▶ doc3: Jones data Jones data Jones data

# Synthetic Example for LDA

- ▶ What should we expect from `model.input`?
  - ▶ doc0: Gilchrist wicket wicket keeper keeper
  - ▶ doc1: wicket captain captain captain keeper keeper
  - ▶ doc2: Jones data architect Gilchrist
  - ▶ doc3: Jones data Jones data Jones data
- ▶ Given number of hidden topics,  $K=2$
- ▶ The documents should be more about one topic than another (based on our prior belief)
- ▶ One topic should be about **sport** and the other about **software** and most words should be more focused on one topic than another
- ▶ Gilchrist may belong to both topics as suggested by data (although we know that Gilchrist was a sportsman)

# Synthetic Example for LDA

## ▶ **compute\_initial\_guesses()**

- ▶ Randomly initializing  $\beta$ \_sufficient-statistics in this case and obtaining row-normalized  $\log(\beta)$  as: [Note:  $\log(\beta) = \{\log(\beta_{i,j})\}$ ]

log beta =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	-1.67219	-1.49886	-2.05742	-2.92477	-3.356	-1.50465	-1.89479
Topic1:	-3.11147	-2.12028	-2.81038	-1.41226	-1.54886	-1.42601	-2.53912

alpha = 0.5

# Synthetic Example for LDA

## ► E-Step begin

**Processing document 0:** Gilchrist(0/0) wicket(1/1) wicket(1/2) keeper(3/3)  
keeper(3/4)

gamma initial = 3 3

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5

gamma =	2.90867	3.09133
Phi =	0.79803	0.20196
	0.63522	0.36477
	0.63477	0.36522
	0.17038	0.82961
	0.17025	0.82974

termID Position

```

doc-e-step() {  $\alpha^{(suffstats)} = 0; \beta^{(suffstats)} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
         $\phi_{n,i}^{t+1} = \beta_{i,w_n} \exp(\psi(\gamma_i^t) - \psi(\sum_{j=1}^K \gamma_j^t))$ ;
      normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
    }
     $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
  curr_ELBO = compute_doc_likelihood();
  converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
  until converged > tolerance or t < Max_T for e-step
  for  $n = 1$  to  $N$ 
    for  $i = 1$  to  $K$ 
       $\beta_{i,w_n}^{(suffstats)} += \phi_{n,i} w_{d,n}$  ;
  for  $i = 1$  to  $K$ 
     $\alpha^{(suffstats)} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
  }
  
```

Beta suff stats =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.79803	1.27	0	0.34063	0	0	0
Topic1:	0.20196	0.7300	0	1.65936	0	0	0

alpha suff stats = -1.56795

# Synthetic Example for LDA

## ► E-Step continue

**Processing document 1:** wicket(1/0) captain(2/1) captain(2/2) captain(2/3)

keeper(3/4) keeper(3/5)

gamma initial = 3.5 3.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;  
0.5 0.5; 0.5 0.5;

gamma = 3.74424 3.25576

Phi = 0.68789 0.31210

0.71509 0.28490

0.71480 0.28519

0.71454 0.28545

0.20602 0.79397

0.20587 0.79412

Beta suff stats =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.798034	1.95789	2.14445	0.752533	0	0	0
Topic1:	0.201966	1.04211	0.855551	3.24747	0	0	0

alpha suff stats = -3.11368

```

doc-e-step() {  $\alpha^{(suffstats)} = 0; \beta^{(suffstats)} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
         $\phi_{n,i}^{t+1} = \beta_{i,w_n} \exp(\psi(\gamma_i^t) - \psi(\sum_{j=1}^K \gamma_j^t))$ ;
      normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
    }
     $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
  curr_ELBO = compute_doc_likelihood();
  converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
  until converged > tolerance or t < Max_T for e-step
  for  $n = 1$  to  $N$ 
    for  $i = 1$  to  $K$ 
       $\beta_{i,w_n}^{(suffstats)} += \phi_{n,i} w_{d,n}$  ;
  for  $i = 1$  to  $K$ 
     $\alpha^{(suffstats)} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
  }
  
```

# Synthetic Example for LDA

## ► E-Step continue

**Processing document 2:** Jones(4/0) data(5/1) architect(6/2) Gilchrist(0/3)

gamma initial = 2.5 2.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;

gamma =	2.80352	2.19648
Phi =	0.18017	0.81982
	0.55360	0.44639
	0.71931	0.28068
	0.85043	0.14957

```

doc-e-step() {  $\alpha^{(suffstats)} = 0; \beta^{(suffstats)} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
         $\phi_{n,i}^{t+1} = \beta_{i,w_n} \exp(\psi(\gamma_i^t) - \psi(\sum_{j=1}^K \gamma_j^t))$ ;
      normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
    }
     $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
  curr_ELBO = compute_doc_likelihood();
  converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
  until converged > tolerance or t < Max_T for e-step
  for  $n = 1$  to  $N$ 
    for  $i = 1$  to  $K$ 
       $\beta_{i,w_n}^{(suffstats)} += \phi_{n,i} w_{d,n}$  ;
  for  $i = 1$  to  $K$ 
     $\alpha^{(suffstats)} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
  }
  
```

Beta suff stats =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	1.6484	1.9578	2.1444	0.7525	0.1801	0.5536	0.7193
Topic1:	0.3515	1.0421	0.8555	3.2474	0.8198	0.4463	0.2806

alpha suff stats = -4.74159

# Synthetic Example for LDA

## ► E-Step continue

**Processing document 3:** Jones(4/0) data(5/1) Jones(4/2) data(5/3) Jones(4/4)  
data(5/5)

gamma initial = 3.5 3.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;  
0.5 0.5; 0.5 0.5

gamma = 0.62764 6.37235

Phi = 0.00661 0.99338

0.03610 0.96389

0.00658 0.99342

0.03594 0.96405

0.00655 0.99344

0.03584 0.96415

Beta suff stats =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	1.6484	1.9578	2.1444	0.7525	0.1999	0.6614	0.7193
Topic1:	0.3515	1.0421	0.8555	3.2474	3.8000	3.3385	0.2806

alpha suff stats = -8.15944

```

doc-e-step() {  $\alpha^{(suffstats)} = 0; \beta^{(suffstats)} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all i and k
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all i
  prev_ELBO = 0; t = 0;
  repeat
    for n = 1 to N {
      for i = 1 to K
         $\phi_{n,i}^{t+1} = \beta_{i,w_n} \exp(\psi(\gamma_i^t) - \psi(\sum_{j=1}^K \gamma_j^t));$ 
      normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
       $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
    }
  curr_ELBO = compute_doc_likelihood();
  converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
  until converged > tolerance or t < Max_T for e-step
  for n = 1 to N
    for i = 1 to K
       $\beta_{i,w_n}^{(suffstats)} += \phi_{n,i} w_{d,n}$  ;
    for i = 1 to K
       $\alpha^{(suffstats)} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
  }
  
```

# Synthetic Example for LDA

## ▶ E-Step end

## ▶ M-Step begin

$$\beta_{ij} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{dni}^* w_{dn}^j$$

Log Beta =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	-1.5900	-1.4180	-1.3270	-2.3742	-3.6997	-2.503	-2.41936
Topic1:	-3.6039	-2.5172	-2.7145	-1.3806	-1.2234	-1.3525	-3.82897

alpha = 0.947154

log\_likelihood= -43.55398 : converged = inf em\_steps = 1 (out of 50)  
TopicHeap0: captain wicket Gilchrist keeper architect data Jones  
TopicHeap1: Jones data keeper wicket captain Gilchrist architect

## ▶ M-Step end

# Synthetic Example for LDA

- ▶ **Iteration 2**
- ▶ **E-Step begin ... E-Step end**
- ▶ **M-Step begin**

$$\beta_{ij} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{dni}^* w_{dn}^j$$

Log Beta =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	-1.7347	-1.3735	-1.3185	-1.6982	-4.6176	-3.4413	-2.55193
Topic1:	-4.0856	-3.3162	-3.7959	-1.6210	-1.0081	-1.0685	-3.99843

alpha = 1.13942

log\_likelihood= -39.944901 : converged = 0.08286 em\_steps = 2 (out of 50)  
TopicHeap0: captain wicket keeper Gilchrist architect data Jones  
TopicHeap1: Jones data keeper wicket captain architect Gilchrist

- ▶ **M-Step end**

# Synthetic Example for LDA

- ▶ **Iteration 3**
- ▶ **E-Step begin ... E-Step end**
- ▶ **M-Step begin**

$$\beta_{ij} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{dni}^* w_{dn}^j$$

Log Beta =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	-1.8527	-1.4130	-1.3982	-1.2863	-5.9505	-4.7532	-2.72934
Topic1:	-4.3027	-4.6651	-5.3798	-2.5688	-0.8214	-0.8397	-3.72119

alpha = 0.885771

log\_likelihood= -37.33781618 : converged = 0.06527 em\_steps = 3 (out of 50)  
TopicHeap0: keeper captain wicket Gilchrist architect data Jones  
TopicHeap1: Jones data keeper wicket captain architect Gilchrist

- ▶ **M-Step end**

# Synthetic Example for LDA

- ▶ **Iteration 4**
- ▶ **E-Step begin ... E-Step end**
- ▶ **M-Step begin**

$$\beta_{ij} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{dni}^* w_{dn}^j$$

Log Beta =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	-1.8896	-1.4260	-1.4240	-1.1591	-7.7516	-6.5435	-2.95167
Topic1:	-4.2712	-6.6943	-7.5649	-4.4922	-0.7617	-0.7649	-3.19709

alpha = 0.605925

log\_likelihood= -34.94676348 : converged = 0.06404 em\_steps = 4 (out of 50)

TopicHeap0: keeper captain wicket Gilchrist architect data Jones

TopicHeap1: Jones data architect Gilchrist keeper wicket captain

- ▶ **M-Step end**

# Synthetic Example for LDA

- ▶ **Iteration 46**
- ▶ **E-Step begin ... E-Step end**
- ▶ **M-Step begin**

$$\beta_{ij} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{dni}^* w_{dn}^j$$

Log Beta =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	-2.3979	-1.2992	-1.2992	-1.0116	-337.72	-337.72	-336.77
Topic1:	-2.3025	-337.14	-337.55	-337.23	-0.9162	-0.9162	-2.3025

alpha = 0.0123227

log\_likelihood= -29.05400653 : converged = 0.00010 em\_steps = 46 (out of 50)

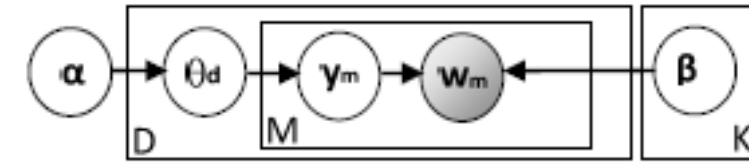
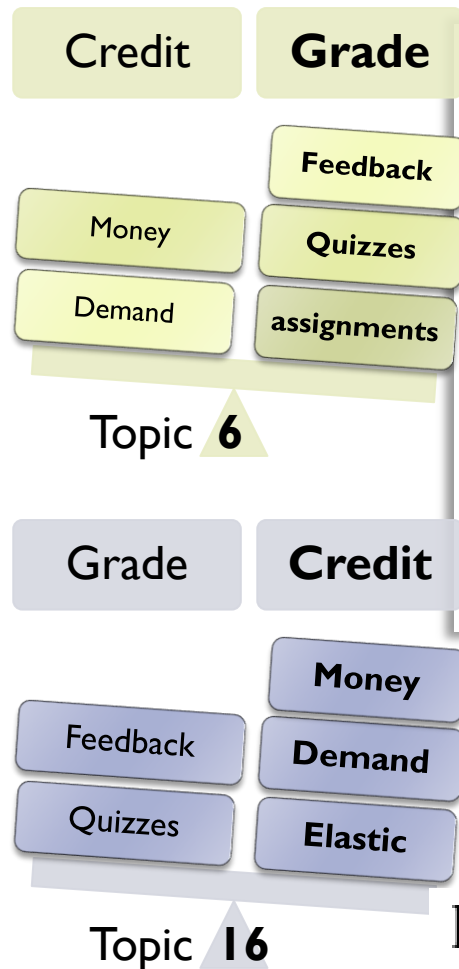
TopicHeap0: keeper wicket captain Gilchrist architect Jones data

TopicHeap1: Jones data Gilchrist architect wicket keeper captain

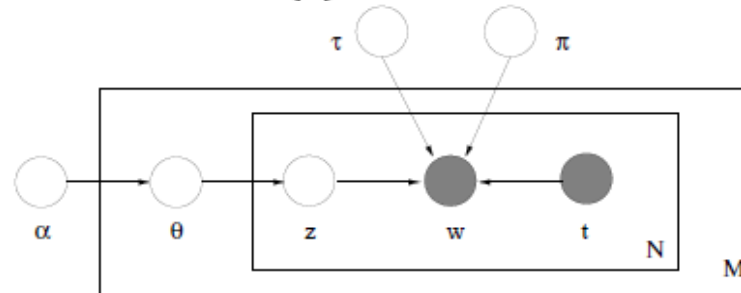
Ties broken arbitrarily

- ▶ **M-Step end**

# LDA and TagLDA



(a) LDA model



(b) TagLDA model

$\beta_{LDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_k$
$\wedge$				$ $
<b>K</b>				$ $
$\vee$				$ $

$\tau_{TagLDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_k$
$\wedge$				$\neq  $
<b>K</b>				$\neq  $
$\vee$				$\neq  $

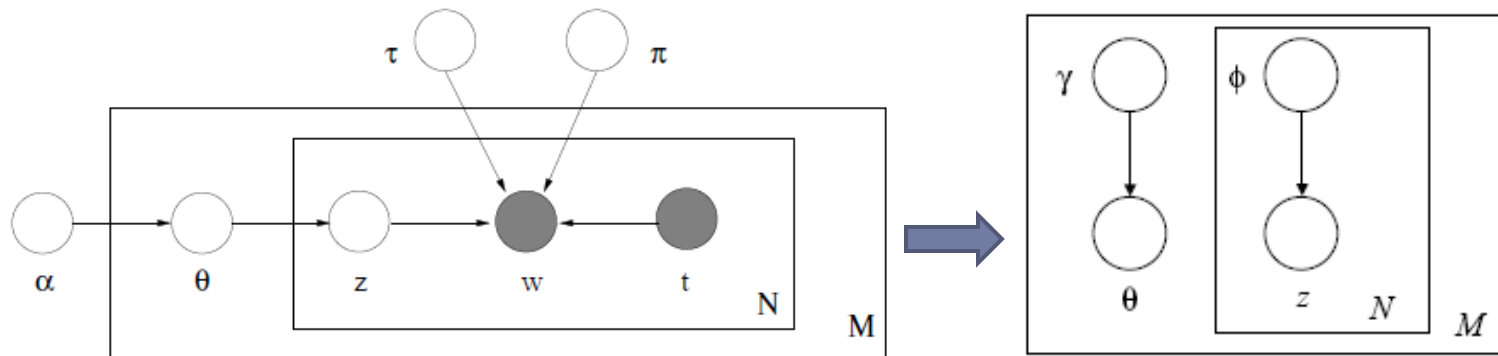
$\pi_{TagLDA}$	$\langle \mathbf{v} \rangle$			$\Sigma_t$
$\wedge$				$\neq  $
<b>T</b>				$\neq  $
$\vee$				$\neq  $

Figure 1: Graphical model representations of LDA and TagLDA models

# Parameters of TagLDA

- ▶  $\tau$  is a set of distributions (topics over words) [Not a multinomial]
- ▶  $\pi$  is a set of distributions (tags over words) [Not a multinomial]
- ▶ For each dataset, we only see their values for the words and tags present in the dataset
- ▶ How much a topic explains the data vs. how much a tag explains the data
- ▶  $\alpha_i$  is a prior for  $\theta_i$  encoding the belief that a document is “about a topic  $i$ .” In basic LDA, all  $\alpha_i$ ’s are equal to  $\alpha$ 
  - ▶ To update  $\tau$  in the m-step() there needs to be TWO  $\tau$ \_marginal-statistics matrices which are updated per document in the e-step
  - ▶ To update  $\pi$  in the m-step() there needs to be TWO  $\pi$ \_marginal-statistics matrices which are updated per document in the e-step
  - ▶  $\alpha$ \_sufficient-statistics similarly is an auxiliary storage for  $\alpha$

# Properties of TagLDA



$$P(w_n = v | z_n, t_n, \tau, \pi) \propto \exp(\tau_{z_n, v} + \tau_{t_n, v})$$

$$p(w|t, \alpha, \tau, \pi) = \int p(\theta|\alpha) \left( \prod_{n=1}^N \sum_{z_n=1}^k p(z_n|\theta) p(w_n|z_n, t_n, \tau, \pi) \right) d\theta$$

$$\begin{aligned} \log p(w|t, \alpha, \tau, \pi) &\geq \int_{\theta} \sum_z q(\theta, z|\gamma, \phi) (\log p(w, \theta, z|t, \alpha, \tau, \pi) - \log q(\theta, z|\gamma, \phi)) d\theta \\ &\equiv L(\gamma, \phi; \alpha, \tau, \pi) \end{aligned}$$

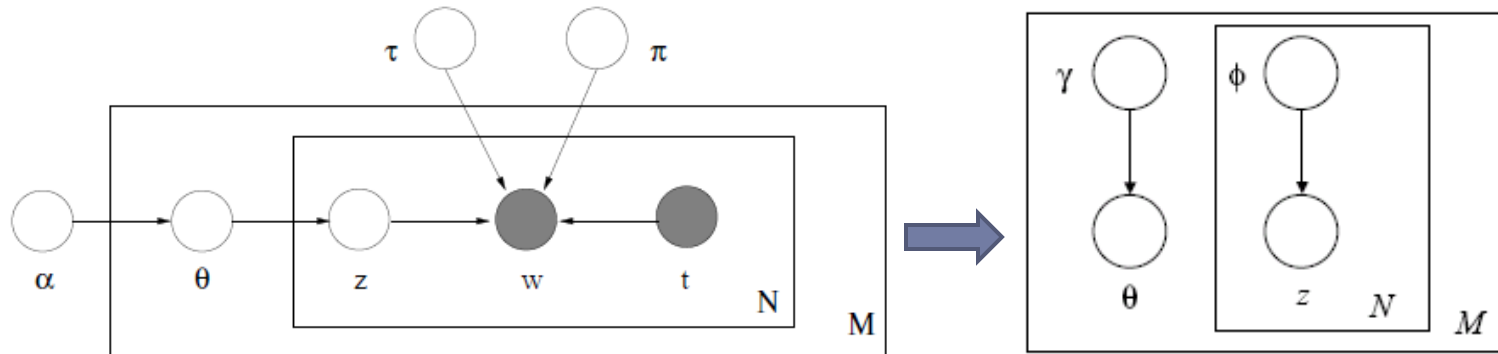
$$L(\gamma, \phi; \alpha, \tau, \pi)$$

$$= E_q[\log p(w, \theta, z|t, \alpha, \tau, \pi)] - E_q[\log q(\theta, z|\gamma, \phi)]$$

$$= E_q[\log p(\theta|\alpha)] + E_q[\log p(z|\theta)] E_q[\log p(w|z, \tau, \pi)]$$

$$- E_q[\log q(\theta)] - E_q[\log q(z)]$$

# Properties of TagLDA



$$E_q[\log p(w|z, \tau, \pi)] = E_q\left[\sum_{n=1}^N \log \frac{\exp(\tau_{z_n, v} + \pi_{t_n, v})}{\sum_{v=1}^V \exp(\tau_{z_n, v} + \pi_{t_n, v})}\right]$$

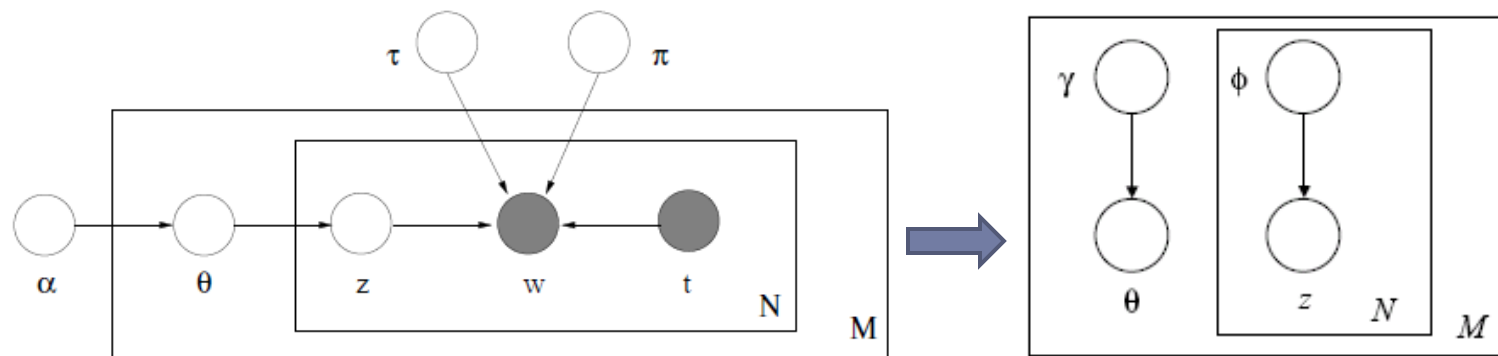
$$\log(x) \leq \zeta^{-1}x + \log(\zeta) - 1, \forall \zeta > 0$$

$$\sum_{n=1}^N E_q\left[\log \sum_{v=1}^V \exp(\tau_{z_n, v} + \pi_{t_n, v})\right] \leq \sum_{n=1}^N \left[ \zeta_n^{-1} \left( \sum_{v=1}^V \sum_{i=1}^k \underbrace{\phi_{ni}}_{E_q(z_n)} \exp(\tau_{z_n, v} + \pi_{t_n, v}) \right) + \log \zeta_n - 1 \right]$$

**NOTE:**  $\tau_{i,v}$  and  $\pi_{t,v}$  are in log-space all throughout TagLDA discussions



# Variational Inference in TagLDA



- ▶ Note:  $\gamma$  (Dirichlet),  $(\phi_1, \phi_2, \dots, \phi_N)$  (Multinomials) and  $\xi$  are *free parameters*

Precompute a  $K \times T$  matrix before every E-step

$$\hat{\xi}_n = \sum_{v=1}^V \sum_{i=1}^k \phi_{ni} \exp(\tau_{i,v} + \pi_{t_n,v})$$

Equ.  
TL- A

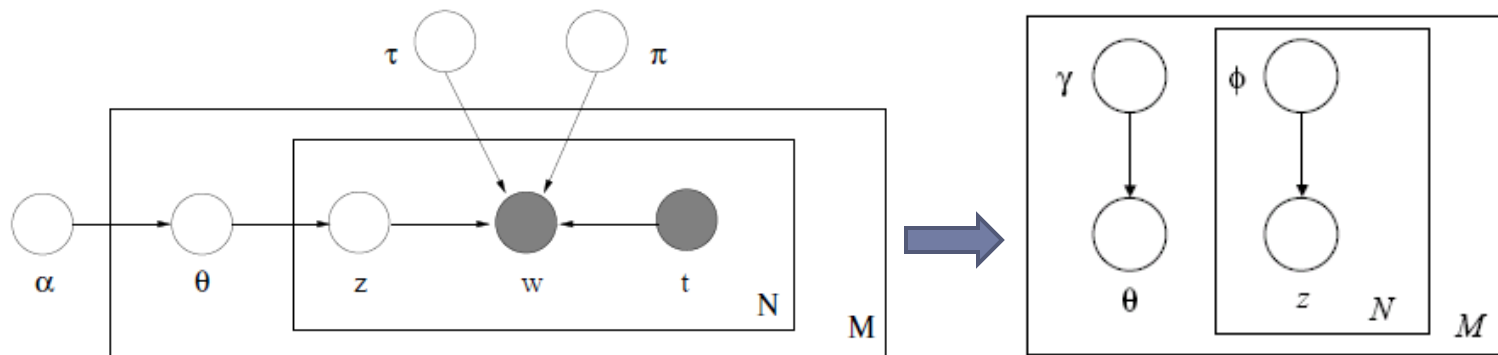
$$\hat{\phi}_{ni} \propto \exp(\psi(\gamma_i) - \psi\left(\sum_{j=1}^k \gamma_j\right) + (\tau_{i,w_n} + \pi_{t_n,w_n}) - \sum_{v=1}^V \xi_n^{-1} \exp(\tau_{i,v} + \pi_{t_n,v}))$$

Equ.  
TL- B

$$\hat{\gamma}_i = \alpha_i + \sum_{n=1}^N \phi_{ni}$$

Equ.  
TL- C

# Variational Inference in TagLDA



$$\tau_{i,v} = \log \left( \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{dn,i} \delta(w_{dn}, v) \right) - \log \left( \sum_{d=1}^M \sum_{n=1}^{N_d} \xi_{dn}^{-1} \phi_{dn,i} \exp(\pi_{t_{dn},v}) \right)$$

Equ.  
TL- D

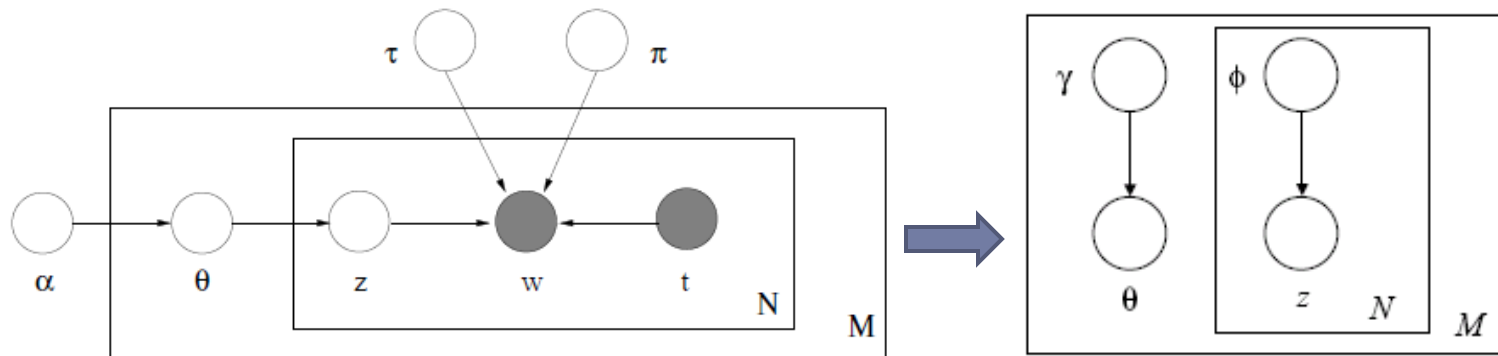
$$\pi_{t,v} = \log \left( \sum_{d=1}^M \sum_{n=1}^{N_d} \sum_{i=1}^k \phi_{dn,i} \delta(t_{dn}, t) \delta(w_{dn}, v) \right) - \log \left( \sum_{d=1}^M \sum_{n=1}^{N_d} \sum_{i=1}^k \xi_{dn}^{-1} \phi_{dn,i} \exp(\tau_{i,v}) \delta(t_{dn}, t) \right)$$

Equ.  
TL- E

$$\alpha_{new} = \alpha_{old} - H(\alpha_{old})^{-1} g(\alpha_{old}) \quad \text{See original paper for } \alpha \text{ optimization}$$

Equ.  
TL- F

# Adding Regularizers in TagLDA



$$\widehat{\mathcal{L}}_{(\cdot)} = \mathcal{L}_{(\cdot)} - \frac{1}{2\sigma^2} \left( \sum_{v=1}^V (\exp(\tau_{i,v}))^2 \right) - \frac{1}{2\sigma^2} \left( \sum_{v=1}^V (\exp(\pi_{t,v}))^2 \right)$$

0-mean,  $\sigma$ -stddev  
Gaussian regularizer for every  $\tau_{i,v}$  and  $\pi_{t,v}$

$$\pi_{t,v} = \log \left( \sum_{d=1}^M \sum_{n=1}^{N_d} \sum_{i=1}^k \phi_{dn,i} \delta(t_{dn}, t) \delta(w_{dn}, v) \right) - \log \left( \sum_{d=1}^M \sum_{n=1}^{N_d} \sum_{i=1}^k \xi_{dn}^{-1} \phi_{dn,i} \exp(\tau_{i,v}) \delta(t_{dn}, t) \right)$$

$$= \log(\text{term}_1^\pi) - \log(\text{term}_2^\pi)$$

letting  $A = \exp(\tau_{i,v})$  or  $\exp(\pi_{t,v})$ ,  $2A = -\sigma^2 \text{term}_2^{(\cdot)} + \sigma \sqrt{\sigma^2 (\text{term}_2^{(\cdot)})^2 + 4 \text{term}_1^{(\cdot)}}$

Update for  $\tau_{i,v}$  is exactly the same with  $\text{term}_1^\tau$  and  $\text{term}_2^\tau$

Equ.  
TL- Er

# Expectation Maximization (EM) for TagLDA

```
▶ compute_initial_guesses() // for beta (corpus distribution)
prev_ELBO = 0;
while (converged > tolerance)
{
    e-step(); // zero initialize 4 marginal statistic matrices and 1
    sufficient statistic for alpha; initialize and update per document
    distributions from corpus distributions; calls doc-e-step() for all
    documents
    m-step(); // update corpus distributions from statistics
    curr_ELBO = compute_likelihood(); // sums up the individual
    doc_likelihoods from compute_doc_likelihood()s in doc-e-step()s
    converged = (prev_ELBO-curr_ELBO)/prev_ELBO;
    prev_ELBO = curr_ELBO;
}
```

# Computing Likelihood for TagLDA

## ► Variational Inference

- We now expand the  $L(\gamma, \phi; \alpha, \tau, \pi)$  into 5 core expressions

$$L(\gamma, \phi; \alpha, \tau, \pi) = \mathbb{E}_q[\log p(\theta|\alpha)] + \mathbb{E}_q[\log p(\mathbf{z}|\theta)] + \mathbb{E}_q[\log p(\mathbf{w}|\mathbf{z}, \tau, \pi)] \\ - \mathbb{E}_q[\log q(\theta)] - \mathbb{E}_q[\log q(\mathbf{z})],$$

- So the likelihood in terms of the model parameters  $(\alpha, \tau, \pi)$  and the variational parameters  $(\xi, \gamma, \phi)$

**compute\_doc\_likelihood()**

$$\log \Gamma\left(\sum_{j=1}^k \alpha_j\right) - \sum_{i=1}^k \log \Gamma(\alpha_i) + \sum_{i=1}^k (\alpha_i - 1)(\psi(\gamma_i) - \psi(\sum_{j=1}^k \gamma_j)) \\ + \sum_{n=1}^N \sum_{i=1}^k \phi_{ni}(\psi(\gamma_i) - \psi(\sum_{j=1}^k \gamma_j)) \\ \sum_{n=1}^N \sum_{i=1}^k \phi_{ni}(\tau_{i,wn} + \pi_{t_n, w_n}) - \sum_{n=1}^N [\zeta_n^{-1} \left( \sum_{v=1}^V \sum_{i=1}^k \phi_{ni} \exp(\tau_{z_n, v} + \pi_{t_n, v}) \right) + \log \zeta_n - 1] \\ - \log \Gamma\left(\sum_{j=1}^k \gamma_j\right) - \sum_{i=1}^k \log \Gamma(\gamma_i) + \sum_{i=1}^k (\gamma_i - 1)(\psi(\gamma_i) - \psi(\sum_{j=1}^k \gamma_j)) \\ - \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} \log \phi_{ni} \equiv L2(\gamma, \phi, \zeta; \alpha, \tau, \pi)$$

# Expectation Maximization (EM) for TagLDA

doc-e-step()  $\{ \alpha^{suffstats} = 0; \tau^{marginalstats_{1,2}} = 0; \pi^{marginalstats_{1,2}} = 0$

initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$

initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$

prev\_ELBO = 0; t = 0;

repeat

for  $n = 1$  to  $N$  {

for  $i = 1$  to  $K$

Update  $\phi_{n,i}^{t+1}$  according to Equ.TL-B;

normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1

$$\gamma_i^{t+1} = \gamma_i^t + \left( \sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t \right)$$

curr\_ELBO = compute\_doc\_likelihood();

converged = (prev\_ELBO - curr\_ELBO)/prev\_ELBO;

until converged > tolerance or t < Max\_T for e-step

for  $n = 1$  to  $N$

Update  $\tau^{marginal_{1,2}}$  according to Equ.TL-Er

for  $i = 1$  to  $K$

Update  $\tau^{marginal_{1,2}}$  according to Equ.TL-Er;

for  $i = 1$  to  $K$

$$\alpha^{suffstats} + = \psi(\gamma_{d,i}) - \psi\left(\sum_{j=1}^K \gamma_{d,j}\right)$$



# Synthetic Example for TagLDA

- ▶ `model.input =`
  - ▶ 5 0:1:0 1:1:1 1:1:1 3:1:1 3:1:1 //wordID:count:tagID
  - ▶ 6 1:1:1 2:1:0 2:1:0 2:1:0 3:1:1 3:1:1 // tagID = 0 means “un-important”
  - ▶ 4 4:1:0 5:1:1 6:1:1 0:1:0 // tagID = 1 means “important”
  - ▶ 6 4:1:0 5:1:1 4:1:0 5:1:1 4:1:0 5:1:1 // count is always 1 in this format
- ▶ `term_tag_to_id.txt =`
  - ▶ 0=un-important            1=important
- ▶ `term_to_id.txt=`
  - ▶ Gilchrist=0 wicket=1 keeper=3 captain=2 Jones=4 data=5  
architect=6
- ▶ `model.input + term_to_id.txt + term_tag_to_id.txt =`
  - ▶ doc0: Gilchrist#0 wicket#1 wicket#1 keeper#1 keeper#1
  - ▶ doc1: wicket#1 captain#0 captain#0 captain#0 keeper#1 keeper#1
  - ▶ doc2: Jones#0 data#1 architect#1 Gilchrist#0
  - ▶ doc3: Jones#0 data#1 Jones#0 data#1 Jones#0 data#1

# Synthetic Example for TagLDA

- ▶ What should we expect from model.input?
  - ▶ doc0: Gilchrist#0 wicket#1 wicket#1 keeper#1 keeper#1
  - ▶ doc1: wicket#1 captain#0 captain#0 captain#0 keeper#1 keeper#1
  - ▶ doc2: Jones#0 data#1 architect#1 Gilchrist#0
  - ▶ doc3: Jones#0 data#1 Jones#0 data#1 Jones#0 data#1
- ▶ Given number of hidden topics,  $K=2$
- ▶ The documents should be more about one topic than another (based on our prior belief)
- ▶ One topic should be about **sport** and the other about **software** and most words should be more focused on one topic than another
- ▶ Gilchrist may belong to both topics as suggested by data (although we know that Gilchrist was a sportsman)

# Synthetic Example for TagLDA

- ▶ What should we expect from model.input?
  - ▶ doc0: Gilchrist#0 wicket#1 wicket#1 keeper#1 keeper#1
  - ▶ doc1: wicket#1 captain#0 captain#0 captain#0 keeper#1 keeper#1
  - ▶ doc2: Jones#0 data#1 architect#1 Gilchrist#0
  - ▶ doc3: Jones#0 data#1 Jones#0 data#1 Jones#0 data#1
- ▶ Given number of observed tags,  $T=2$
- ▶ We know wicket, keeper, data, architect are important words (Word level domain knowledge)
- ▶ If Gilchrist were to be “important” based on the dominance of important words on his topic would he be more or less “important” than Jones?
- ▶ Jones should remain less important because his topic is conditioned on “importance” less than that of Gilchrist’s

# Synthetic Example for TagLDA

## ▶ **compute\_initial\_guesses()**

- ▶ Randomly initializing  $\tau_{\text{marginal}}$ -statistics in this case and obtaining  $\log(\tau)$  as: [Note:  $\log(\tau) = \{\log(\tau_{i,j})\}$ ]
- ▶ Similarly for  $\log(\pi)$

Log tau =

	Gilchrist wicket	captain	keeper	Jones	data	architect	
Topic0:	0.1189	0.3112	-1.2743	0.4446	0.2119	0.3041	0.5725
Topic1:	0.1181	0.5153	-0.0278	0.2744	0.3097	0.0091	-0.4654

Log pi =

	Gilchrist wicket	captain	keeper	Jones	data	architect	
Tag0:	-1.67219	-1.49886	-2.05742	-2.92477	-3.356	-1.50465	-1.89479
Tag1:	-3.11147	-2.12028	-2.81038	-1.41226	-1.54886	-1.42601	-2.53912

alpha = 0.5

# Synthetic Example for TagLDA

## ► E-Step begin

**Processing document 0:** Gilchrist(0/0/0) wicket(1/1/1) wicket(1/2/1) keeper(3/3/1)

keeper(3/4/1)

gamma initial = 3 3

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5

gamma =	1.54437	4.45563
Phi =	0.20887	0.79112
	0.18003	0.81997
	0.17886	0.82113
	0.23915	0.76084
	0.23743	0.76256

Tau suff stats 1=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.2088	0.3588	0	0.4765	0	0	0
Topic1:	0.7911	1.6411	0	1.52341	0	0	0

Tau suff stats 2=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.0122	0.0456	0	0.0290	0	0	0
Topic1:	0.0463	0.2087	0	0.0928	0	0	0

termID Position Tag

```

doc-e-step() {
   $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,j}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
        normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
      }
       $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
    curr_ELBO = compute_doc_likelihood();
    converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
    until converged > tolerance or t < Max_T for e-step
  for  $n = 1$  to  $N$ 
    Update  $\pi^{marginal,2}$  according to Equ. TL-Er
  for  $i = 1$  to  $K$ 
    Update  $\tau^{marginal,2}$  according to Equ. TL-Er;
  for  $i = 1$  to  $K$ 
     $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
  }
  
```

# Synthetic Example for TagLDA

## ► E-Step begin

**Processing document 0:** Gilchrist(0/0/0) wicket(1/1/1) wicket(1/2/1) keeper(3/3/1)

keeper(3/4/1)

gamma initial = 3 3

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5

gamma =	1.54437	4.45563
Phi =	0.20887	0.79112
	0.18003	0.81997
	0.17886	0.82113
	0.23915	0.76084
	0.23743	0.76256

Pi suff stats 1=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	1	0	0	0	0	0	0
Tag1:	0	2	0	2	0	0	0

Pi suff stats 2=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.1133	0	0	0	0	0	0
Tag1:	0	0.4076	0	0.3450	0	0	0

termID Position Tag

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$

initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$

initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$

prev\_ELBO = 0; t = 0;

repeat

for  $n = 1$  to  $N$  {

for  $i = 1$  to  $K$

Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;

normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1

$\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$

curr\_ELBO = compute\_doc\_likelihood();

converged = (prev\_ELBO - curr\_ELBO)/prev\_ELBO;

until converged > tolerance or t < Max\_T for e-step

for  $n = 1$  to  $N$

Update  $\pi^{marginal,2}$  according to Equ. TL-Er

for  $i = 1$  to  $K$

Update  $\tau^{marginal,2}$  according to Equ. TL-Er;

for  $i = 1$  to  $K$

$\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$



► 47 alpha suff stats = -1.95729

# Synthetic Example for TagLDA

## ► E-Step continue

**Processing document 1:** wicket(1/0/1) captain(2/1/0) captain(2/2/0) captain(2/3/0)

keeper(3/4/1) keeper(3/5/1)

gamma initial = 3.5 3.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;  
0.5 0.5; 0.5 0.5;

gamma = 0.67885 6.32115

Phi = 0.03696 0.96303

0.01283 0.98716

0.01280 0.98719

0.01276 0.98723

0.05197 0.94802

0.05151 0.94848

Tau suff stats 1=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.2088	0.3958	0.0384	0.5800	0	0	0
Topic1:	0.7911	2.6041	2.9616	3.4199	0	0	0

Tau suff stats 2=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.0122	0.0503	0.0060	0.0354	0	0	0
48Topic1:	0.0463	0.3320	0.4657	0.2094	0	0	0

```

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,1} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
        normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
      }
       $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
    curr_ELBO = compute_doc_likelihood();
    converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
    until converged > tolerance or t < Max_T for e-step
    for  $n = 1$  to  $N$ 
      Update  $\pi^{marginal,1,2}$  according to Equ. TL-Er
      for  $i = 1$  to  $K$ 
        Update  $\tau^{marginal,1,2}$  according to Equ. TL-Er;
      for  $i = 1$  to  $K$ 
         $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
    }
  }
  
```



# Synthetic Example for TagLDA

## ► E-Step continue

**Processing document 1:** wicket(1/0/1) captain(2/1/0) captain(2/2/0) captain(2/3/0)  
 keeper(3/4/1) keeper(3/5/1)  
 gamma initial = 3.5 3.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;  
 0.5 0.5; 0.5 0.5;

gamma =	0.67885	6.32115
Phi =	0.03696	0.96303
	0.01283	0.98716
	0.01280	0.98719
	0.01276	0.98723
	0.05197	0.94802
	0.05151	0.94848

Pi suff stats 1=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	1	0	3	0	0	0	0
Tag1:	0	3	0	4	0	0	0

Pi suff stats 2=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.1133	0	0.2957	0	0	0	0
49Tag1:	0	0.6183	0	0.6815	0	0	0

alpha suff stats = -5.22157

```

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
        normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
      }
       $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
    curr_ELBO = compute_doc_likelihood();
    converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
    until converged > tolerance or t < Max_T for e-step
    for  $n = 1$  to  $N$ 
      Update  $\pi^{marginal,2}$  according to Equ. TL-Er
      for  $i = 1$  to  $K$ 
        Update  $\tau^{marginal,2}$  according to Equ. TL-Er;
      for  $i = 1$  to  $K$ 
         $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
    }
  }
  
```



# Synthetic Example for TagLDA

## ► E-Step continue

**Processing document 2:** Jones(4/0/0) data(5/1/1) architect(6/2/1) Gilchrist(0/3/0)

gamma initial = 2.5 2.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;

gamma =	4.09844	0.90155
Phi =	0.86101	0.13898
	0.90549	0.09450
	0.95316	0.04683
	0.87875	0.12124



```

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
      normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
    }
     $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
  curr_ELBO = compute_doc_likelihood();
  converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
  until converged > tolerance or t < Max_T for e-step
  for  $n = 1$  to  $N$ 
    Update  $\pi^{marginal_{1,2}}$  according to Equ. TL-Er
    for  $i = 1$  to  $K$ 
      Update  $\tau^{marginal_{1,2}}$  according to Equ. TL-Er;
  for  $i = 1$  to  $K$ 
     $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
  }
  
```

Tau suff stats 1=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	1.0876	0.3958	0.0384	0.5800	0.8610	0.9054	0.9531
Topic1:	0.9123	2.6041	2.9616	3.4199	0.1389	0.0945	0.0468

Tau suff stats 2=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.0875	0.0503	0.0060	0.0354	0.0953	0.1961	0.08919

► 50Topic1: 0.0567 0.3320 0.4657 0.2094 0.0153 0.0204 0.00438

# Synthetic Example for TagLDA

## ► E-Step continue

**Processing document 2:** Jones(4/0/0) data(5/1/1) architect(6/2/1) Gilchrist(0/3/0)

gamma initial = 2.5 2.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;

gamma =	4.09844	0.90155
Phi =	0.86101	0.13898
	0.90549	0.09450
	0.95316	0.04683
	0.87875	0.12124



```

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
      normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
    }
     $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
  curr_ELBO = compute_doc_likelihood();
  converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
  until converged > tolerance or t < Max_T for e-step
  for  $n = 1$  to  $N$ 
    Update  $\pi^{marginal,2}$  according to Equ. TL-Er
    for  $i = 1$  to  $K$ 
      Update  $\tau^{marginal,2}$  according to Equ. TL-Er;
  for  $i = 1$  to  $K$ 
     $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
}
  
```

Pi suff stats 1=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	1	0	3	0	1	0	0
Tag1:	1	3	0	4	0	1	1

Pi suff stats 2=

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.1133	0	0.2957	0	0.1205	0	0
Tag1:	0.1374	0.6183	0	0.6815	0	0.1612	0.2091

alpha suff stats = -7.70207



# Synthetic Example for TagLDA

## ► E-Step continue

**Processing document 3:** Jones(4/0/0) data(5/1/1) Jones(4/2/0) data(5/3/1) Jones(4/4/0)  
data(5/5/1)

gamma initial = 3.5 3.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;  
0.5 0.5; 0.5 0.5

gamma =	4.6866	2.3134
Phi =	0.65012	0.34987
	0.74024	0.25975
	0.65295	0.34704
	0.74261	0.25738
	0.65572	0.34427
	0.74493	0.25506

Tau suff stats I =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	1.0876	0.3958	0.0384	0.5800	2.8198	3.1333	0.9531
Topic1:	0.9123	2.6041	2.9616	3.4199	1.1801	0.8667	0.0468

Tau suff stats I =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.0875	0.0503	0.0060	0.0354	0.3155	0.6823	0.08919

► 52Topic1: 0.0567 0.3320 0.4657 0.2094 0.1324 0.1890 0.00438

```

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
        normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
      }
       $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
    curr_ELBO = compute_doc_likelihood();
    converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
    until converged > tolerance or t < Max_T for e-step
    for  $n = 1$  to  $N$ 
      Update  $\pi^{marginal_{1,2}}$  according to Equ. TL-Er
      for  $i = 1$  to  $K$ 
        Update  $\tau^{marginal_{1,2}}$  according to Equ. TL-Er;
      for  $i = 1$  to  $K$ 
         $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
    }
  }
  
```

# Synthetic Example for TagLDA

## ► E-Step continue

**Processing document 3:** Jones(4/0/0) data(5/1/1) Jones(4/2/0) data(5/3/1) Jones(4/4/0)  
data(5/5/1)

gamma initial = 3.5 3.5

Phi initial = 0.5 0.5; 0.5 0.5; 0.5 0.5; 0.5 0.5;  
0.5 0.5; 0.5 0.5

gamma =	4.6866	2.3134
Phi =	0.65012	0.34987
	0.74024	0.25975
	0.65295	0.34704
	0.74261	0.25738
	0.65572	0.34427
	0.74493	0.25506

Pi suff stats l =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	1	0	3	0	4	0	0
Tag1:	1	3	0	4	0	4	1

Pi suff stats l =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.1133	0	0.2957	0	0.4953	0	0

► 53Tag1: 0.1374 0.6183 0 0.6815 0 0.6278 0.2091

alpha suff stats = -9.40611

```

doc-e-step() {  $\alpha^{suffstats} = 0; \tau^{marginalstats,2} = 0; \pi^{marginalstats,2} = 0$ 
  initialize  $\phi_{n,i}^0 = 1/K$  for all  $i$  and  $k$ 
  initialize  $\gamma_i^0 = \alpha_i + N/K$  for all  $i$ 
  prev_ELBO = 0; t = 0;
  repeat
    for  $n = 1$  to  $N$  {
      for  $i = 1$  to  $K$ 
        Update  $\phi_{n,i}^{t+1}$  according to Equ. TL-B;
        normalize  $\phi_{n,i}^{t+1}$  so that row-sum=1
      }
       $\gamma_i^{t+1} = \gamma_i^t + (\sum_{n=1}^N \phi_{n,i}^{t+1} - \sum_{n=1}^N \phi_{n,i}^t)$ 
    curr_ELBO = compute_doc_likelihood();
    converged = (prev_ELBO - curr_ELBO)/prev_ELBO;
    until converged > tolerance or t < Max_T for e-step
    for  $n = 1$  to  $N$ 
      Update  $\pi^{marginal,2}$  according to Equ. TL-Er
      for  $i = 1$  to  $K$ 
        Update  $\tau^{marginal,2}$  according to Equ. TL-Er;
      for  $i = 1$  to  $K$ 
         $\alpha^{suffstats} += \psi(\gamma_{d,i}) - \psi(\sum_{j=1}^K \gamma_{d,j})$ 
    }
  }
  
```

# Synthetic Example for TagLDA

## ▶ E-Step end

## ▶ M-Step begin

Log Tau =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.6932	0.1897	-0.9520	0.3976	1.1176	1.072	0.6235
Topic1:	0.6176	1.0689	1.1010	1.2513	0.7150	0.5202	-0.8475

Log Pi =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.6364	-100	1.15719	-100	1.26278	-100	-100
Tag1:	0.6244	1.0648	-100	1.2167	-100	1.2299	0.5887

alpha = 0.678869

log\_likelihood = -44.89721451 : converged = inf em\_steps = 1 (out of 50)

TopicHeap0: Jones data Gilchrist architect keeper wicket captain  
TopicHeap1: keeper captain wicket Jones Gilchrist data architect  
TagHeap0: Jones captain Gilchrist wicket keeper data architect  
TagHeap1: data keeper wicket Gilchrist architect captain Jones

## ▶ M-Step end

# Synthetic Example for TagLDA

## ▶ Iteration 2

## ▶ E-Step begin ... E-Step end / M-Step begin

Log Tau =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.6885	-0.2695	-1.0752	-0.1760	1.1497	1.2356	0.6473
Topic1:	0.6108	1.13506	1.1025	1.2530	-0.2226	-0.2289	-1.2173

Log Pi =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.6452	-100	1.1153	-100	1.19934	-100	-100
Tag1:	0.6549	1.1609	-100	1.2732	-100	1.2753	0.6575

alpha = 0.52364

log\_likelihood= -26.50176340 : converged = 0.40972 em\_steps = 2 (out of 50)

TopicHeap0: data Jones Gilchrist architect keeper wicket captain

TopicHeap1: keeper wicket captain Gilchrist Jones data architect

TagHeap0: Jones captain Gilchrist wicket keeper data architect

TagHeap1: data keeper wicket architect Gilchrist captain Jones

## ▶ M-Step end

# Synthetic Example for TagLDA

## ▶ Iteration 3

## ▶ E-Step begin ... E-Step end / M-Step begin

Log Tau =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.6587	-0.9884	-1.5983	-0.8960	1.1676	1.2421	0.6516
Topic1:	0.6275	1.1430	1.0697	1.2602	-0.9764	-1.0936	-1.7940

Log Pi =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.6329	-100	1.0739	-100	1.1829	-100	-100
Tag1:	0.6541	1.1516	-100	1.2682	-100	1.2514	0.6556

alpha = 0.35718

log\_likelihood= -23.45330186 : converged = 0.11503 em\_steps = 3 (out of 50)

TopicHeap0: data Jones Gilchrist architect keeper wicket captain

TopicHeap1: keeper wicket captain Gilchrist Jones data architect

TagHeap0: Jones captain Gilchrist wicket keeper data architect

TagHeap1: keeper data wicket architect Gilchrist captain Jones

## ▶ M-Step end

# Synthetic Example for TagLDA

## ▶ Iteration 4

## ▶ E-Step begin ... E-Step end / M-Step begin

Log Tau =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.6495	-1.9324	-2.4608	-1.8273	1.1657	1.228	0.6491
Topic1:	0.6228	1.1423	1.0526	1.2568	-1.8753	-2.1385	-2.7003

Log Pi =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.6228	-100	1.0537	-100	1.1699	-100	-100
Tag1:	0.6497	1.1441	-100	1.2587	-100	1.2307	0.6499

alpha = 0.247851

log\_likelihood= -21.21241265 : converged = 0.09555 em\_steps = 4 (out of 50)

TopicHeap0:	data	Jones	Gilchrist	architect	keeper	wicket	captain
TopicHeap1:	keeper	wicket	captain	Gilchrist	Jones	data	architect
TagHeap0:	Jones	captain	Gilchrist	wicket	keeper	data	architect
TagHeap1:	keeper	data	wicket	architect	Gilchrist	captain	Jones

## ▶ M-Step end

# Synthetic Example for TagLDA

## ▶ Iteration 5

## ▶ E-Step begin ... E-Step end / M-Step begin

Log Tau =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.6453	-3.0959	-3.5984	-2.9767	1.1604	1.2163	0.6453
Topic1:	0.6168	1.1392	1.0399	1.2528	-2.9986	-3.3860	-3.9017

Log Pi =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.6167	-100	1.0402	-100	1.1614	-100	-100
Tag1:	0.6454	1.1395	-100	1.2531	-100	1.2167	0.6454

alpha = 0.181814

log\_likelihood= -19.68089582 : converged = 0.07220 em\_steps = 5 (out of 50)

TopicHeap0: data Jones architect Gilchrist keeper wicket captain  
TopicHeap1: keeper wicket captain Gilchrist Jones data architect  
TagHeap0: Jones captain Gilchrist wicket keeper data architect  
TagHeap1: keeper data wicket Gilchrist architect captain Jones

## ▶ M-Step end

# Synthetic Example for TagLDA

## ▶ Iteration 27

## ▶ E-Step begin ... E-Step end / M-Step begin

Log Tau =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Topic0:	0.6414	-44.2895	-44.8044	-44.1473	1.1544	1.2054	0.6414
Topic1:	0.6110	1.1363	1.0281	1.2492	-44.1279	-44.7341	-45.2408

Log Pi =

	Gilchrist	wicket	captain	keeper	Jones	data	architect
Tag0:	0.6110	-100	1.0281	-100	1.1544	-100	-100
Tag1:	0.6414	1.1363	-100	1.2492	-100	1.2054	0.6414

alpha = 0.020894

log\_likelihood = -17.49778259 : converged = 0.00047 em\_steps = 27 (out of 50)

TopicHeap0:	data	Jones	Gilchrist	architect	keeper	wicket	captain
TopicHeap1:	keeper	wicket	captain	Gilchrist	Jones	data	architect
TagHeap0:	Jones	captain	Gilchrist	wicket	keeper	data	architect
TagHeap1:	keeper	data	wicket	Gilchrist	architect	captain	Jones

## ▶ M-Step end

# References

- ▶ David M. Blei, Andrew Y. Ng, and Michael I. Jordan. Latent dirichlet allocation. *JMLR*, 3:993–1022, 2003
- ▶ Xiaojin Zhu, David Blei, and John Lafferty. Taglda: Bringing document structure knowledge into topic models. *UWisc Technical Report TR-1533*, 2006