Introduction

Knowledge of the size and distribution of human population is essential for understanding and responding to many social, political, economical, and environmental problems (Liu et al. 2006). In the United States, the decennial census is the primary source of demographic data. Although the U.S. Census Bureau conducts census surveys at the household level, the data are released only by aggregated enumeration units because of confidentiality issues and administrative purposes. Researchers may encounter two problems when using aggregated census data to conduct research.

The first problem is associated with cartographic visualization of population density. Choropleth mapping of population density creates the impression that population density is uniform within census units. This may be acceptable when the mapping area consists of a large number of units and the map is generated to visualize the overall trend of population variation. However, when the mapping area only consists of a few census units there may be a need to further delineate different population density zones within the units for better visualization of population distribution. Specifically, a census unit may contain inhabited and uninhabited areas, residential and non-residential land use, and areas of different population densities, particularly in suburban or rural areas where population is sparse. Therefore, for the purpose of visualizing within-unit population distribution, in contrast to visualizing the differences between units, population density maps should, whenever possible, try to map where people actually live. Accordingly, methods are needed to delineate different population density zones within census units.

The second problem for researchers using census population data is that they may need to estimate population counts for areas not coinciding with boundaries of census units. For example, a subdivision for a proposed redevelopment project does not always have the same boundaries as do the census units, yet developers may need to estimate how many people live in the subdivision for cost and benefit analysis. Watersheds generally do not have the same boundaries as the census units, yet researchers may need to estimate watershed populations for environmental impact assessments. City planners may need to estimate populations within a half mile buffer of a proposed railroad route or a proposed landfill site to estimate how many local residents would be affected by the...
projects. All these examples show that we need methods for estimating populations of arbitrary boundaries.

This study presents a dasymetric mapping method for intra-census-unit population mapping and estimation. We disaggregated census block populations based on their land-use information and image-texture statistics of the semi-variance.


According to Wu et al. (2005), past studies relevant to population estimation in GIS and remote sensing can be grouped into areal interpolation and statistical modeling, depending on the intended goal and the required information. Areal interpolation studies use census population data as the input and apply certain interpolation techniques to obtain refined population data, usually for the purpose of transforming population data from one set of spatial units to another. In contrast, statistical modeling studies infer a statistical relationship between population and other physical and socio-economic variables, often for the purpose of estimating intercensal populations for urban areas or populations of areas where it is difficult to conduct the census. Nevertheless, some researchers also incorporate statistical modeling into areal interpolation of population (e.g., Harvey 2002b; Yuan et al. 1997).

Areal interpolation of population may be further separated into two categories, depending on whether the interpolation is based on mathematical functions or ancillary information (Wu et al. 2005). For areal interpolation based on mathematical functions there are point-based methods and areal-based methods, depending on whether the population data for computational input are in digital form of points or areas. Examples of point-based interpolation include Martin (1989), Bracken (1991), and Martin (1996). Examples of area-based interpolation include Tobler (1979) and Rase (2001). Areal-based methods usually have the volume-preserving property, i.e., the summation of interpolated population data to individual census units is the same as the original census unit populations.

Population is related to other information such as land use, transportation, and topography. Ancillary information relevant to population distribution, therefore, can be used to assist areal interpolation of population. Areal interpolation using ancillary information is referred to as the dasymetric mapping method, and this approach is generally volume preserving (Wu et al. 2005).

The most commonly used ancillary information for population interpolation is land use/land cover (LULC) data (e.g., Yuan et al. 1997; Mennis 2003; Holt et al. 2004). Others include topographic data (Wright 1936), election districts demographic data (e.g., Flowerdew and Green 1989; 1991), road network data (e.g., Xie 1995; Reibel and Bufalino 2005; Hawley and Moellering 2005), and remote sensing image spectral and textural statistics (Harvey 2002b).

Land use/land cover-based dasymetric mapping assumes the same population densities for the same LULC classes and redistributes census populations to land use areas. The population densities for different land-use classes can be determined from sampling (e.g., Mennis 2003), from regression analysis (e.g., Yuan et al. 1997; Langford et al. 1991), or based on the domain knowledge of researchers (e.g., Eicher and Brewer 2001). Regression analysis seems to provide a more objective approach since population density measures are derived based on the entire land-use and population dataset.

Land use/land cover data are usually derived from remote sensing images through digital image classification. Due to the difficulty of classifying detailed categories of urban land use from remote sensing images (Wu et al. 2006), remote sensing-derived LULC data only provide general categories of urban land use (e.g., residential vs. non-residential) as the ancillary information for population interpolation. Compounded with the errors associated with digital image classification, LULC-based dasymetric mapping for urban areas always has a degree of error. Nevertheless, if detailed and accurate urban land-use data are available, e.g., data of single family, multi-family, industrial, and commercial land uses, then population interpolation based on the data should be more reliable. Such land-use data are most likely to be obtained through image interpretation and manual digitizing. They are usually in vector format with a high degree of precision.

Another category of literatures relevant to population estimation is statistical modeling. Depending on the scale of analysis, researchers have used five types of predictor variables in regression analyses of population (Wu et al. 2005), including urban areas (e.g., Tobler 1969; Lo and Welch 1977; Prosperie and Eyton 2000), land-use areas (e.g., Kraus et al. 1974; Weber 1994; Lo 2003), classi-
fied dwelling units count (e.g., Hsu 1971; Lo and Chan 1980; Lo 1989), image pixel statistics (e.g., Webster 1996; Harvey 2002a; Liu et al. 2006), and other physical or socioeconomic characteristics (e.g., Green and Monier 1959; Dobson et al. 2000; Liu and Clarke 2002).

We separated population estimation research in GIS and remote sensing into two categories: areal interpolation and statistical modeling. This notwithstanding, statistical modeling approaches have been used for areal interpolation of population (specifically the dasymetric mapping method) in some studies, in the context of estimating population densities for land-use classes (e.g., Yuan et al. 1997; Langford et al. 1991; Flowerdew and Green 1989) or estimating population by image pixels (e.g., Harvey 2002b).

Although LULC-based dasymetric mapping is an improvement over traditional choropleth mapping, important variations of population distribution within individual land-use classes may still be ignored. In contrast, Harvey (2002b) disaggregated census populations by pixels using a variety of spectral and textural statistics of remote sensing images. This pixel-based dasymetric mapping allows the mapping of detailed variations of population distribution. Its limitation is that the relationship between population and combinations of image pixel statistics cannot be interpreted. In addition there is the possibility that non-residential land areas have the same image pixel statistics as residential land areas. As a result, assuming people only live on residential land, this complexity would make modeling of population by image pixel statistics less reliable. This study combines the strength of LULC-based and pixel-based dasymetric mapping by modeling population from image pixel statistics for different land uses. In this approach, we can separate residential versus non-residential areas as well as areas of different residential land use classes. Then we can model detailed variations of population distribution from image pixel statistics. The volume-preserving strength of dasymetric mapping method is maintained.

Past studies have indicated a strong relationship between population density and image spectral and texture statistics (Lo 1995; Webster 1996; Harvey 2002a). Texture statistics from remotely sensed images measure the degree of spectral variation between pixels and, therefore, indicate the degree of landscape heterogeneity for a given area. For example, residential areas with small houses and small distances between houses represent more heterogeneous landscape and usually have higher values of image texture statistics, such as variance, standard deviation, and entropy. In contrast, residential areas with large houses and long distances between houses represent more homogeneous landscape and usually have lower values of image texture statistics (Bian and Xie 2004; Liu et al. 2006; Wu et al. 2006).

In this study, we used the texture statistic of semi-variance to correlate with population densities. Semi-variance statistics for an area are calculated as half the average of the squared difference between paired pixel values separated at a certain distance, called the lag. The mathematical function of semi-variance, γ(h), separated by a lag, h, can be expressed as:

\[
\gamma(h) = \frac{1}{2N_h} \sum_{i=1}^{N} (z_i - z_{i+h})^2
\]

where:

- \(N_h\) = the number of paired pixels separated by lag \(h\); and
- \(z_i, z_{i+h}\) = values of a pair of pixels separated by lag \(h\).

A graph of semi-variance against the lag is called the variogram. The variogram is usually fitted with a mathematical model to indicate the extent of spatial autocorrelation across space, or the degree of landscape heterogeneity at certain lag-scales. As the lag increases, the semi-variance generally becomes larger. At a certain distance, called the range, the semi-variance reaches its limit, called the sill. The range represents the limit of spatial dependence and indicates the distance over which the values would be similar (Burrough and McDonnell 1998).

In this study, semi-variance is used to indicate the degree of landscape heterogeneity, in a way to that of other image texture statistics. The advantage of the semi-variance is its ability to indicate landscape heterogeneity at different scales. Common texture statistics, such as variance and standard deviation, only give us an average measure across scales. We expect population density would better correlate to image texture statistics at certain scales.

**A Case Study of Austin, Texas**

The City of Austin, the capital of Texas, provides a suitable environment to explore the proposed census population interpolation method. With a land area of approximately 650 square kilometers in 2005, Austin is not too large and has a variety of land use types. Between 1990 and 2000, the city’s population has grown from 465,622 to 656,562, or 41 percent, with an
Using block-level census data is more appropriate than using other higher-level census data, e.g., block groups and tracts data, because a census block is the smallest census enumeration unit and block-level statistics are suitable for inferring sub-block-level populations.

In the second stage of our analysis, we applied the derived regression model to a study site and estimated population densities for individual land-use areas within blocks. We then calculated and rescaled their populations to maintain the summed populations to blocks.

The Census 2000 block-level population data for the Austin area were obtained from the U.S. Census Bureau’s American FactFinder website (U.S. Census Bureau 2000). Digital aerial photographs and land-use data of the year 2000 were obtained from the City of Austin Neighborhood Planning and Zoning Department (NPZD). The digital orthophotos have a spatial resolution of about 0.61 meters (two feet) and green, red, and near-infrared spectral bands. These high-spatial-resolution images allowed us to identify structures on the ground and have an understanding of how population density related to housing and land-use patterns. When calculating image semi-variance statistics, we resampled the digital orthophotos to approximately five meters (16 feet) of spatial resolution so as to achieve computational efficiency. Compared to Système Probatoire d’Observation de la Terre (SPOT) or Landsat Enhanced Thematic Mapper Plus (ETM+) data used in other dasy-metric studies, the fine resolution images we used allowed us to detect detailed spectral variation between human structures and vegetation through image texture statistics in a complicated urban landscape, and proved to better model population densities.

The year 2000 land-use data are in a vector polygon format and have 13 categories (Table 1). The NPZD generated the data based on the geography of Travis Central Appraisal District (TCAD) tax parcel polygons and land-use information from different sources. Digital infrared and panchromatic aerial photos were the primary sources used to determine land use through visual interpretation (COA 2000). Other sources of land-use information include historical land-use data, the TCAD appraisal record, and the land development record. The NPZD also performed quality control measures by field checks. The land-use data are used to select residential land-use blocks for building population density models by different residential land-use types. As the land-use data are at the tax parcel level, generally much smaller than census blocks, the data were considered accurate for the purposes of this study.

There are not many large lot single-family and mobile homes land-use areas in Austin; these two categories were thus combined with the single-family land-use category. We only created a single-family land-use model and a multi-family land-use model for inferring block-level population densities.

To build regression models between block population density and image statistics, we first selected 120 and 100 sample census blocks that are entirely within, respectively, single- and multi-family land use. The sample blocks represent various housing patterns for the same land use in the Austin area. The population densities for individual sample blocks were then calculated by dividing block populations with block areas. We calculated the image texture statistics of semi-variance for sample blocks. A visual basic program was written in the ArcGIS® ArcObjects (Burke 2003) programming environment to calculate the semi-variance of paired image pixel values that were separated at lag distance, based on pixel size within individual sample block areas. We used the near-infrared band of the digital orthophotos to calculate semi-variance statistics because a preliminary analysis indicates that population density is more related to the semi-variance statistics from the infrared band than from other bands. The infrared band shows more contrast between vegetation and human structures, which is relevant to variations of population densities.

We calculated the semi-variance at lag 2 pixels to lag 50 pixels for sample blocks and generated corresponding variograms to investigate which scales of semi-variance to use for population modeling. Figure 1 shows the variograms for 120 single-family sample blocks, and Figure 2 shows the variograms for 100 multi-family sample blocks. One can see that, at the same lag, single-family blocks generally have greater semi-variances (note the difference
in vertical scales in the figures) than multi-family blocks. This means that single-family land-use areas have more heterogeneous landscape patterns than multi-family land-use areas and, thus, have greater semi-variance texture statistics. Notice that each block has a different variogram pattern. As no single lag is particularly significant for differentiating blocks, we tested semi-variance at all lags in statistical modeling of block population densities.

Our response variable is block population density, and predictor variables are semi-variances at different lags. We first conducted an exploratory analysis examining histograms of all variables one at a time for both the single-family land use model and the multi-family land use model. If it was highly skewed, a logarithmic transformation was applied. The logarithmic transformation had the effect of reducing asymmetry (Chatterjee et al. 2000). The histograms of block population densities (Figure 3) and all semi-variance variables

<table>
<thead>
<tr>
<th>Code</th>
<th>Major Group</th>
<th>Included Uses</th>
</tr>
</thead>
<tbody>
<tr>
<td>50</td>
<td>Large Lot</td>
<td>Single-family homes on lots greater than 10 acres</td>
</tr>
<tr>
<td>100</td>
<td>Single-family</td>
<td>Single-family detached, Two-Family Attached</td>
</tr>
<tr>
<td>113</td>
<td>Mobile homes</td>
<td>Mobile homes</td>
</tr>
<tr>
<td>200</td>
<td>Multi-family</td>
<td>Three/Fourplex, apartment/condo, group quarters, retirement</td>
</tr>
<tr>
<td>300</td>
<td>Commercial</td>
<td>Retail and general merchandise, apparel and accessories, furniture and home furnishings, grocery and food sales, eating and drinking, auto related, entertainment, personal services, lodgings, building services</td>
</tr>
<tr>
<td>400</td>
<td>Office</td>
<td>Administrative offices, financial services (banks), medical offices, research and development</td>
</tr>
<tr>
<td>500</td>
<td>Industrial</td>
<td>Manufacturing, warehousing, equipment sales and service, recycling and scrap, animal handling</td>
</tr>
<tr>
<td>560</td>
<td>Mining</td>
<td>Resource extraction, quarries</td>
</tr>
<tr>
<td>600</td>
<td>Civic</td>
<td>Semi-institutional housing, hospital, government services, educational facilities, meeting and assembly facilities, cemeteries, day care facilities</td>
</tr>
<tr>
<td>700</td>
<td>Open Space</td>
<td>Parks, recreational facilities, golf courses, reserves and protected areas, water drainage areas and detention ponds</td>
</tr>
<tr>
<td>800</td>
<td>Transportation</td>
<td>Railroad facilities, transportation terminal, aviation facilities, parking facilities, right-of-way and traffic islands</td>
</tr>
<tr>
<td>860</td>
<td>Right-of-way</td>
<td>Right-of-way and traffic islands</td>
</tr>
<tr>
<td>870</td>
<td>Utilities</td>
<td>Utility services, radio towers, communication service facilities, water/wastewater facilities</td>
</tr>
<tr>
<td>900</td>
<td>Undeveloped/Rural</td>
<td>Rural uses, vacant land, land under construction</td>
</tr>
<tr>
<td>940</td>
<td>Water</td>
<td>Inundated areas such as lakes and rivers where delineated</td>
</tr>
<tr>
<td>999</td>
<td>Unknown</td>
<td>a lot requiring further information to determine how it is used</td>
</tr>
</tbody>
</table>

Table 1. Year 2000 land-use data of the City of Austin.

Figure 1. Variograms for single-family census blocks.
in both land use models are not highly skewed and were not transformed.

We further examined pairwise scatter plots to explore the relationships between block population densities and semi-variance variables, as well as between the semi-variance variables. Figure 4 shows that semi-variances at different lags are highly correlated, which suggested a multi-collinearity problem. Nevertheless, we tested all semi-variances in regression analysis in order to investigate how population densities are related to semi-variances at different scales and to build a model with the best predictive power.

Figure 4 also shows that population densities do not have linear relationships with individual semi-variance variables. However, we cannot conclude that there is no statistical relationship between population densities and combinations of semi-variance variables. In addition, due to the simplicity of linear regression models and the ease with which they can be interpreted, we hypothesized a linear relationship between population densities and combinations of semi-variance variables. We then conducted a standard statistical regression analysis.

We used the backward elimination procedure to find appropriate combinations of variables. In the procedure, the regression with the full variable set is calculated first, and insignificant variables are removed in turn. In this way, the multi-collinearity problem can be handled better than with the forward selection procedure and the stepwise method. The final models from backward elimination retained seven semi-variances as the predictor variables for the single-family model and eight semi-variances for the multi-family model:
PDS = 1185 + 2.1*S1 – 2.4*S2 + 1.9*S3
– 0.6*S4 + 1.0*S5 – 1.1*S6 + 0.1*S7 (2)

where:
PDS = block population densities for the single-family land use model; and
M1 to M8 = semi-variances at lag 2, lag 6, lag 8, lag 14, lag 26, lag 28, and lag 46, respectively.

PDM = 0.5 – 3.6*M1 + 6.0*M2 – 4.0*M3 + 2.7*M4
– 2.7*M5 + 2.3*M6 – 4.7*M7 + 2.8*M8 (3)

where:
PDM = block population densities for the multi-family land use model; and
M1 to M8 = semi-variances at lag 2, lag 6, lag 8, lag 26, lag 28, lag 44, lag 48, and lag 50, respectively.

The fit indices, $R^2$, were 0.72 and 0.67 for single-family and multi-family land use models, respectively, indicating that approximately 72 percent and 67 percent of the variability in block population densities can be explained by combinations of semi-variance variables. The significance values of the F statistic in the ANOVA analysis of variance test were both less than 0.05, indicating that the variation explained by the models was not due to chance.

These results confirm that using a range of scales of semi-variance to model population densities is preferable to using semi-variances at similar scales. In other words, because population densities are related to landscape heterogeneity at different scales, they are better modeled by different scales of semi-variance. The single-family land-use model predicted population densities better at smaller lags than did the multi-family land-use model. This is probably because single-family land use has smaller building sizes and is more sensitive to spatial statistics at smaller scales. The single-family land-use model also had a higher $R^2$, which indicates that there is a more regular and predictable relationship between landscape heterogeneity and population densities in single-family land use than in multi-family land use.

We further verified the standard regression assumptions of linearity, normality of errors, zero mean of errors, constant variance of errors, and independence of errors. Graphical methods were used for the tests because of their simplicity and ease of interpretation. The histogram of the standardized residual (Figure 5), the normal probability plot of the standardized residual (Figure 6), and the scatter plot of the standardized residual versus the standardized predicted value (Figure 7) all show that the assumptions were satisfied.

We investigated whether spatial autocorrelation exists in block-level population density, because serious spatial autocorrelation would violate the independence assumptions underlying regression analysis (O’Sullivan and Unwin 2003; Longley et al. 2005). Moran’s I statistic for 7077 blocks in the Austin area was calculated using ArcGIS (Mitchell 2005) in order to determine whether the pattern of block population density is clustered, dispersed, or random. A Moran’s I value near +1.0 indicates
clustering, a value near –1.0 indicates dispersion, and a value close to zero indicates a random pattern. The corresponding z score indicates whether the clustering or dispersion could be the result of random chance or is statistically significant. For example, at a significance level of 0.05, a z score would have to be less than –1.96 or greater than 1.96 to be statistically significant.

We obtained a Moran’s I of 0.01 and a z score of 110. Consequently, we concluded that such a low spatial autocorrelation of population density at the block-level would not have important effect on the regression analysis. The low spatial autocorrelation of population density may be caused by the fact that, in the Austin area, current residential land-use areas are confined to small tracts of land and scattered among other land use.

A test area of 251 blocks was selected to apply the population density models for population disaggregation (Figure 8). This area has a variety of land use types, and many blocks contain both residential and non-residential land use. To redistribute block populations to single-family and multi-family land-use areas, we first overlaid land-use data with census block data to derive individual land-use areas within blocks. Then, digital orthophotos were overlaid to calculate semi-variance statistics for the within-block land-use areas. After that, regression models were applied to estimate population densities for

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**Figure 5.** Histograms of standardized residual for (a) single-family and (b) multi-family land-use models.

**Figure 6.** Normal probability plots of the standardized residual for (a) single-family and (b) multi-family land-use models.
those areas. If the estimated population density was negative, it was adjusted to zero, given that population density cannot be negative (e.g., Lo 1995; Harvey 2002a).

We calculated estimated populations for the within-block land-use areas by multiplying their respective areas with their estimated population densities. We then summed the areal populations to census blocks and compared the summed totals with original census block populations to derive their ratios. Finally, population estimates for within-block land-use areas were rescaled to maintain original census block populations.

To visualize the population disaggregation results, a graduated color thematic map of population densities by small land-use areas within blocks was generated (Figure 9). The map revealed detailed variation of population densities within blocks, in contrast to the traditional population density map by census blocks (Figure 10). A single-family land-use area of the map (Figure 11) shows that the map distinguishes three population density zones related to the spatial patterns of houses.

To quantitatively assess the proposed disaggregation approach and compare it with the traditional LULC-based population disaggregation, we performed a population disaggregation test that maintains census block-group populations instead of block populations. The sum of the block populations thus obtained can be further compared with census block populations for the purpose of accuracy assessment.

The mean absolute relative error (MARE) may be used to compare estimated block populations with original census block populations. The MARE for the 251 blocks in our study was calculated as follows:

$$MARE = \frac{1}{m} \sum_{i=0}^{m} \left| \frac{P_i - Y_i}{Y_i} \right| \times 100\% \quad (4)$$
where:
\( P_i \) = the estimated population for the \( i \)th census block;
\( Y_i \) = the census reported population for the \( i \)th census block; and
\( m \) = the number of census blocks under investigation.

The MARE gives an overall estimate of the percentage of original block populations that were under- or over-estimated. This measure was adopted because it is easy to interpret. The MARE for the proposed population disaggregation approach is approximately 11.8 percent, indicating that on average, approximately 12 percent of the original block populations are either over- or under-estimated.

To compare the proposed disaggregation approach with the traditional LULC-based population disaggregation, we first calculated the average population densities for single-family land use at 2.94 persons per 1000 square meters and for multi-family land use at 10.40 persons per 1000 square meters. Then we estimated populations for within-block land-use areas based on population density. The areal populations were further rescaled to maintain census block-group populations. Lastly, The MARE was calculated to assess the extent to which the aggregated block populations deviate from census block populations.

The MARE was approximately 19.2 percent, which is much higher than the MARE obtained with our disaggregation method. We therefore concluded that initial land-use stratification and further texture statistical modeling has a higher overall accuracy than the traditional LULC-based population disaggregation. Modeling population densities by image semi-variance statistics without land-use stratification had a MARE of approximately 14.6 percent, further indicating the advantage of our method.

We investigated the distribution of estimation errors and whether the errors are spatially correlated by calculating the relative error for an \( i \)th census block \( (RE_i) \):

\[
RE_i = \frac{P_i - Y_i}{Y_i} \times 100\%
\]

where:
\( Y_i \) = the census population for the \( i \)th census block;
Pi = the estimated population for the ith census block;
m = the number of census blocks under investigation; and
REi = the relative error for the ith census block.

The relative errors for the 251 blocks ranged from -9.8 percent to 98.2 percent. They are skewed to the right (Figure 12), indicating that most block populations were over-estimated. The cause of this is not clear. The relative errors of block population estimates mapped in Figure 13 do not show whether the errors are spatially correlated. The 0.03 Moran’s I statistic we calculated, and the corresponding z score of 6, both indicate that the errors were not spatially autocorrelated.

**Discussion**

We only built a single-family and a multi-family land-use model for the Austin area. More levels of model stratification may be needed for other urban areas in order to model subtypes of residential land use that are significant to those area, e.g., mobile home land use and large-lot single-family land use.

Regression models derived in this study may not be directly applicable to other urban areas, because housing geography and the corresponding image texture statistics vary considerably among U.S. cities. A more appropriate strategy would thus be to develop a new population density model for the area of interest, using the multi-lag semi-variance statistics and land-use stratification approach we proposed.

A limitation is that this method requires detailed urban land-use data for model stratification. The data quality/accuracy issues that may affect population density estimation using our method include miscounting of the census data on houses under construction or otherwise unoccupied, time difference between census and image acquisition, and outdated land-use data. However, since our final population disaggregation by small within-block land-use areas is constrained by census
block population totals, all errors are also constrained and would not have a significant impact on population estimation for large areas that cover numerous blocks (Fisher and Langford 1996).

## Conclusion

In this paper we present an improved dasymetric mapping method for remodeling census populations. The method models areal population densities from texture statistics of remote sensing images within the same land-use stratification, while maintaining census block population totals, as is the case in common dasymetric mapping approaches. The proposed approach combines the strength of LULC- and pixel-based dasymetric mapping. The results show improved accuracies compared to either LULC- or pixel-based approaches using the same parameters. The proposed population estimation method may be applied to estimate intercensal populations in conjunction with the analysis of remote sensing data taken during the current year.

Land use/land cover stratification combined with pixel-based dasymetric mapping allows reliable population estimation at large scales, particularly for urban areas. The refined population maps provide a more accurate representation of population distribution than conventional maps of population density by census blocks or land-use types. The disaggregated populations may allow for population estimation within arbitrary boundaries, particularly when integrating with other spatial data of different spatial units, such as watersheds or ecological zones.

Future research may incorporate other population-relevant variables into the regression models, such as building heights and other socio-economic statistics. However, for the models to be practical, their parameters should be readily available or easily extractable from remote sensing images. The presented method utilizes image texture statistics and existing land-use data to remodel census population, which suggests that it is feasible for fine-scale population estimation.

## REFERENCES


