

MULTI-INDEX MODELS

Fall 2000

- 1. Assumptions that residuals are uncorrelated from a single-index model turns out to be inaccurate. This assumption can be improved by assuming additional indexes.**
- 2. Once again multi-index models have developed a life of their own beyond their original purpose of estimating covariances.**

Consider two index model:

$$r_{it} = a_i + b_{i1} I_{1t} + b_{i2} I_{2t} + e_{it}$$

Where:

$$(1). E \begin{pmatrix} I_{1t} & I_{2t} \end{pmatrix} = 0$$

$$(2). E \begin{pmatrix} e_{it} \end{pmatrix} = 0$$

$$(3). E \begin{pmatrix} e_{it} & I_{1t} & \text{or} & I_{2t} \end{pmatrix} = 0$$

$$(4). E \begin{pmatrix} e_{it} & e_{jt} \end{pmatrix} = 0$$

$$(5). a_i, b_{i1} \text{ and } b_{i2} \text{ are constant.}$$

only real assumption is $E \begin{pmatrix} e_{it} & e_{jt} \end{pmatrix} = 0$

Mean return:

$$E\left(r_{it}\right) = a_i + b_{i1} \bar{I}_1 + b_{i2} \bar{I}_2$$

$$\text{Var}\left(r_{it}\right) = b_{i1}^2 s_1^2 + b_{i2}^2 s_2^2 + s_{ei}^2$$

in other words:

$$s_{ei}^2 = b_{i2}^2 s_2^2 + s_{ei}^2$$

$$\text{cov}\left(r_i, r_j\right) = b_{i1} b_{j1} s_1^2 + b_{i2} b_{j2} s_2^2$$

Note:

1. if add $E\begin{pmatrix} e_i & e_j \end{pmatrix}$ get historical.

2. assuming $E\begin{pmatrix} e_i & e_j \end{pmatrix} = b_{i2} b_{j2} \frac{s^2}{I^2} + 0$

Generalizing to more than two indexes is straight forward.

Types of indexes:

(1). Statistically derived

(2). Portfolios of securities

A. S&P, H-L, B-M

B. S&P, H-L, B-M, Bonds

Fama French

E&G

(3). Economic Factors

A. Δ IP, surprise in inflation

B. Surprise GNP

(4). Market and Industry

A. S&P

B. Industry factors

In bond area:

$$R_{it} = \bar{R}_i + OAS - D_i \frac{\Delta r}{1+r} - V_i dV_i + C_i (\text{spread}) + e_i$$

Issues:

(1). Real influences

(2). Parsimonious

Uses:

(1). Covariance structure

(2). Selection of exposure

(3). Return attribution

(4). Portfolio evaluation

